

# How English domiciled graduate earnings vary with gender, institution attended, subject and socio-economic background

### IFS Working Paper W16/06

Jack Britton Lorraine Dearden Neil Shephard Anna Vignoles



## How English domiciled graduate earnings vary with gender, institution attended, subject and socioeconomic background\*

JACK BRITTON
Institute for Fiscal Studies, London
jack\_b@ifs.org.uk

LORRAINE DEARDEN

Institute for Fiscal Studies, London and Institute of Education, University College London lorraine\_d@ifs.org.uk

NEIL SHEPHARD

Department for Economics and Department of Statistics, Harvard University

shephard@fas.harvard.edu

ANNA VIGNOLES

Department of Education, University of Cambridge
av404@cam.ac.uk

April 13, 2016

#### Abstract

This paper uses tax and student loan administrative data to measure how the earnings of English graduates around 10 years into the labour market vary with gender, institution attended, subject and socioeconomic background. The English system is competitive to enter, with some universities demanding very high entrance grades. Students specialise early, nominating their subject before they enter higher education (HE). We find subjects like Medicine, Economics, Law, Maths and Business deliver substantial premiums over typical graduates, while disappointingly, Creative Arts delivers earnings which are roughly typical of non-graduates. Considerable variation in earnings is observed across different institutions. Much of this is explained by student background and subject mix. Based on a simple measure of parental income, we see that students from higher income families have median earnings which are around 25% more than those from lower income families. Once we control for institution attended and subject chosen, this premium falls to around 10%.

<sup>\*</sup>Many civil servants and policy makers have helped us gain access to the data which is the core of this paper. Although it is difficult to pick out a small group who helped most, we must thank in particular Daniele Bega, Dave Cartwright, Nick Hillman, Tim Leunig and David Willetts who were all crucial in making this project happen. Of course we solely are responsible for any errors. We are particularly grateful to the Nuffield Foundation for their financial support. The Nuffield Foundation is an endowed charitable trust that aims to improve social well-being in the widest sense. It funds research and innovation in education and social policy and also works to build capacity in education, science and social science research. The Nuffield Foundation has funded this project, but the views expressed are those of the authors and not necessarily those of the Foundation. HM Revenue & Customs (HMRC) and Student Loans Company (SLC) have agreed that the figures and descriptions of results in the attached document may be published. This does not imply HMRC's or SLC's acceptance of the validity of the methods used to obtain these figures, or of any analysis of the results. Copyright of the statistical results may not be assigned. This work contains statistical data from HMRC which is Crown Copyright and statistical data from SLC which is protected by Copyright, the ownership of which is retained by SLC. The research datasets used may not exactly reproduce HMRC or SLC aggregates. The use of HMRC or SLC statistical data in this work does not imply the endorsement of either HMRC or SLC in relation to the interpretation or analysis of the information.

**Keywords**: Administrative data; Graduate earnings; Human capital; Labour earnings; Degree subject; Student loans.

#### 1 Introduction

This paper uses high quality administrative tax data to provide estimates of English graduates' earnings and shows how they vary according to gender, subject studied, institution attended and socioeconomic background of the student. Over and above improving our understanding of the sources of variation in graduates' earnings, this paper can also inform issues relating to social mobility. In England it is well known that access to university, on average, varies substantially by the level of parental income and that students from poorer families access different types of universities than those from wealthier backgrounds. However, the question of whether graduates' earnings vary according to their socioeconomic background amongst graduates attending similar universities and taking the same subject has remained poorly understood, thus far limited by data availability. Our unique administrative database offers substantial advantages in addressing this crucial question. The findings are also relevant for myriad other issues that benefit from better information on variation in graduates' earnings, including: student choice of subject and institution; better information for schools to help advise and guide students whilst at school; and the operation and cost of the higher education finance system.

#### 1.1 The core of our paper

This paper is the first to use administrative tax records to provide estimates of sub-populations of English domiciled graduates' labour earnings over their early life course. This new and abundant longitudinal data source allows us to study how graduates' earnings develop and vary according to the subject they studied, the institution they attended, their gender and an indicator of their family's income status. We record means and quantiles of these sub-populations. We also have data from the UK's Higher Education Statistics Agency (HESA) on the socioeconomic background and pre-HE academic achievement of the students studying the same subject in the same institution (these data are not linked to our earnings data at individual level). With these data we can estimate a model that controls for student intake and hence enables us to report a value added measure of the degree.

The core of our paper bores into a unique database developed and documented by Britton et al. (2015) which was built through the hard linking of anonymised individual level data from the English student loan book, a book owned by the UK's Secretary of State for Business, Innovation and Skills and operated by the Student Loan Company (SLC), with the corresponding income

tax records held by Her Majesty's Revenue and Customs (HMRC), the UK tax authority. The book starts with new borrowers from 1998 and we look in detail at tax records for the tax years 2008/09 to 2012/13. Throughout, we refer to these borrowers as graduates, though in practice we do not observe whether they complete their course. We have several cohorts of students, from 1998 through to 2011 (though information on the latter cohorts is limited as many are still enrolled on their courses during the years we observe), and primarily focus on the 1999-2005 cohorts. This allows us to follow graduates through their most crucial career developing years. Hence another contribution of this paper is to provide some insight into graduates' earnings growth some years into graduates' careers, rather than just on entry into the labour market as is the case for much of the UK literature.

Britton et al. (2015) defined the population of English domiciled student loan borrowers and constructed the labour earnings variables for this database, as well as providing English wide summary statistics of earnings (e.g. median earnings in the tax year 2011/12 of English former borrowers who started HE in 1999 and are female).

For the first time we use these administrative data to characterise the properties of earnings for sub-populations of borrowers (graduates) and shows how they vary by gender, degree subject and higher education institution. An example of this is that we can observe the 2011/12 earnings of female Law students at Manchester University who started at that university in 2001 (we call this the 2001 cohort). We think about these sub-populations in a number of ways, by looking at means and quantiles of:

- Unconditional earnings.
- Predicted earnings for this sub-population using an individual level parental income indicator and HESA profiles of the socioeconomic background and pre-HE academic achievement of the students in each subject/institution combination.

This second approach can be used to provide a conditional estimate of the earnings of graduates from different institutions or taking different degree subjects, after controlling for differences in some key characteristics of the individual or the institution and is our approximation of a value-added measure of the university by subject. We are mindful however, that selection into degree courses will mean that our estimates are not going to tell us about the causal impact of a particular degree on earnings. Further we do not have detailed information about the education achievement or other characteristics beyond gender and age of non-graduates and hence, whilst we can compare graduate earnings to non-graduate earnings, we cannot calculate a formal rate of return on a particular degree. Instead we focus on measures of variation in graduates' earnings that are themselves of

considerable value.

The data we use and the analysis it generates is highly original. Whilst other UK surveys such as the Labour Force Survey (LFS) and the Destinations of Leavers from Higher Education (DLHE) survey have information on subject of study and institution, the latter has only recently been collected by the LFS, limiting the sample sizes available to researchers. A current source of UK data on graduate earnings by subject and institution is the DLHE survey, which looks at full time equivalent earnings 3 years past graduation. However, our administrative data offers major advantages in terms of scale, quality and duration. We will see that it is important to model both institution and subject. Further, Britton et al. (2015) suggested that much of the interesting HE impacts on earnings emerge after the 3 year horizon when postgraduate education and initial career training has largely been completed. Certainly, the diversity of graduate earnings takes a while to emerge, but they are on stark display after ten years when the earnings of some graduate sub-populations are rapidly accelerating. For this reason, it is essential we observe a longer time span of gradates' earnings than is provided in the DLHE and our administrative data offers an opportunity to do so.

In the rest of this section we will explain why our work matters and place our contribution in the context of the existing literature. We also provide an outline for the rest of the paper.

#### 1.2 Public policy

Policy change has been rapid in the English higher education system in recent years, with major reforms to the finance system, a lifting of the cap on student numbers and some quite dramatic changes in the composition of the student body with a large increase in the number of EU students accessing UK higher education Dorling (2016). Against this backdrop of rapid change there is a need to improve our understanding of the diversity of the sector and the variability in graduates' earning outcomes. Of course students go to university for many reasons other than for pecuniary gain and many graduates do socially valuable jobs that are not necessarily higher paying. Nonetheless, reliable information on graduates' earnings is crucially important from a public policy perspective.

There are three principal reasons for us to better understand the diversity of graduates' earnings (given in alphabetical order rather than order of importance): (i) Funding; (ii) Information; (iii) Social mobility.

The UK Government runs an income contingent loan system for funding English domiciled students at UK higher education providers (HEPs). As a result, the cost to the Government depends crucially upon the path of earnings of graduates over the first 30 years of their career (e.g. Browne (2010), Barr (2007), Barr and Shephard (2010)). Our data enable us to measure earnings after the first decade of a graduates' career providing additional insights into this issue.

Students from all socioeconomic backgrounds need to be informed about their options. One aspect of this, which the UK Government tries to provide using the six month and 3 year DHLE data, is earnings data (see the unistats.direct.gov.uk website which provides the Key Information Set (KIS) as a summary of crucial information for students for each degree course). Unfortunately, the DHLE data misses most of the acceleration in graduate earnings which some, but by no means all, graduate careers see (acceleration depends upon institution and subject choice). Hence the current information available to students strongly under reports the diversity of graduate earnings across subject and institutional choices. This is likely to be more damaging for students who come from families and communities who are less informed about potential HE choices. This paper will consider whether the administrative data analysed here can start to bridge this information gap.

A central focus for this paper is what the diversity of graduates' earnings observed in our data may mean for social mobility, defined as the relationship between parental background (measured by an indicator of family income in this case) and a child's eventual labour market success (measured by their income up to ten years after graduation). Of course social immobility has many causes, but much of the focus has been on the problems of low achievement at school amongst poorer students and the relatively lower likelihood of more disadvantaged students accessing HE and in particular high status universities. In this paper we ask whether students from poorer backgrounds who attend similar universities and study the same subject end up earning less in the labour market than their more advantaged counterparts? If students from poorer backgrounds appear to earn less for a given degree choice, there may be implications for firms with regards to their hiring policies (e.g. the role of unpaid internships) and universities in relation to the career guidance and support they give their students. Whilst we are unable to estimate the causal impact of family income on graduates' earnings<sup>1</sup> we are able to provide a full description of the extent to which the earnings of graduates from different socioeconomic backgrounds appear to be equalised (or not) after leaving higher education.

#### 1.3 Academic literature

This work will contribute to an important literature on the impact of higher education on individuals' earnings and human capital (Blundell et al. (2005), Becker (1962)). Estimating the causal impact of education on earnings is challenging, due problems with ability bias driving degree choice and the difficulty in separating the productivity value of education from its signalling value (Card (1999, 2012)). In the absence of experimental or quasi-experimental data, like much of the existing literature, we can only provide descriptions of the variation in earnings across different sub

<sup>&</sup>lt;sup>1</sup>Since we are unable to control for all factors that may influence earnings and that are correlated with family income, such as the level of non-cognitive skill of a graduate.

populations of graduates. However, given our additional controls for the socioeconomic profile and pre-HE achievement of students taking different degrees we can take some account of potential ability biases.

The work is innovative in that, although the empirical literature from the UK has already shown substantial variation in graduate earnings that has increased over time (Blundell et al. (2005), Bratti et al. (2005), Chevalier (2011), Hussain et al. (2009), Sloane and O'Leary (2005), Smith and Naylor (2001), Walker and Zhu (2011)), researchers have not thus far been able to assess adequately how graduate earnings vary according to the university attended. Theoretically we would expect that different institutions may add different amounts of human capital value and hence influence students' success in the labour market. This work will also complement existing research on the variation in earnings by subject of degree (Sloane and O'Leary (2005), Walker and Zhu (2011)). Walker and Zhu (2013) have recently built on their earlier work which used the LFS to explore differences in graduates' earnings by subject (Walker and Zhu (2011)) to investigate how lifetime graduate earnings vary by both subject and institution type (as measured by broad university groupings, for example the "Russell Group", and "Millionplus"). However, given that there is as yet no usable survey data that contains both degree subject and institution for graduates of different ages, and since they do not have access to the administrative data we use here, they had to simulate the earnings profiles by splicing different survey data sets together (similar to work carried out by the Institute for Fiscal Studies on this issue, e.g. Chowdry et al. (2013)). This work has suggested greater private and social returns to a degree than many had previously estimated. Specifically, the private lifetime earnings return was estimated by Walker and Zhu to be in the order of £168k for men and £252k for women, with the social benefits exceeding the private benefit in both cases. Their more recent analysis confirmed their earlier work and showed substantial differences in private returns by degree subject. By contrast they found insignificant differences in returns by institution type. However, they acknowledged that with the data they had available they were unable to test the robustness of these findings.

Estimates of the variation in graduates' earnings are likely to be somewhat country specific because the degree of subject specialisation and institutional hierarchy varies across countries. Hence we restrict our review of empirical evidence largely to the UK. However, it is important to note that existing evidence points to increasing heterogeneity in graduates' earnings in the US, linked to both choice of college (Monks (2000)) and college major (Arcidiacono (2004)).

This paper will also contribute to understanding about how graduates' outcomes vary by their socioeconomic background. We have already noted that UK students from poor backgrounds are far less likely to attend university in the first place, and they are particularly less likely to

attend a higher status university. Whilst most of the difference in access to HE by socioeconomic background is explained by differences in rich and poor students' prior achievement, there remains a small socioeconomic gap in HE participation conditional on prior achievement (Chowdry et al. (2012), Croxford and Raffe (2013)). Irrespective of the reason, if students from poor backgrounds are far less likely to access the kinds of degree courses associated with very high earnings, this will affect their life chances and hence is crucial from a policy perspective.

Over and above differential access to different types of HE, individuals' socioeconomic background may also continue to have an effect on their labour market outcomes after graduation. This might be because students from more advantaged backgrounds have higher levels of (non-cognitive) skills (see for example Blanden et al. (2007)) skills that are not measured by their highest education level, or by their degree subject or institution. Alternatively, advantaged graduates may earn more because they have greater levels of social capital and are able to use their networks to secure higher paid employment. The literature on this is quite limited in the UK but does suggest that graduates from more advantaged backgrounds, particularly privately educated students, achieve higher status occupations and earn a higher return to their degree (Bukodi and Goldthorpe (2011b), Bukodi and Goldthorpe (2011a), Macmillan et al. (2013), Crawford and Vignoles (2014)). For example, Crawford and Vignoles (2014) indicate that graduates who attended private secondary schools earn around 7% more per year, on average, than state school students 3.5 years after graduation, even when comparing otherwise similar graduates and allowing for differences in degree subject, university attended and degree classification. Bratti et al. (2005) use the British Cohort Study (BCS) which follows a cohort born in 1970 and found little evidence of variation in the return to a degree by social class. Dolton and Vignoles (2000) found that the earnings return for graduates varied according to whether the individual attended a private school or a state school. This work was based on a cohort of 1980 graduates and the private school wage premium for graduates was 7% for males and 0% for females, conditional on subject of degree and institution. The suggestion that private school graduates earn an additional premium over and above the return to their degree is also supported by evidence from Naylor (2002) for a cohort of 1993 graduates (3\% wage premium) and by Green et al. (2012) using the National Child Development Study 1958 cohort and the 1970 BCS referred to earlier. They found more generally that the private school wage premium increased from 4% for the earlier cohort to 10% for the later one, a finding which held for graduates. Whilst we do not have an indicator of whether or not the individual attended a private school in our data. we do have an indicator of parental income and can therefore explore the variation in earnings by parental income level, for a given degree.

Another innovation of this paper is the fact that, for the first time, we have been able to use

administrative tax records for graduates some considerable time into their careers (up to 11 years after graduation) to produce very high quality estimates of earnings. The paper will therefore also improve understanding of the specific administrative databases being analysed and contribute to the literature on the use of administrative data, and its advantages and disadvantages, as discussed in Savage and Burrows (2009), Webber (2009) and Card et al. (2010). There is a limited but growing literature on the application of large scale administrative data to understand the outcomes from education; a review can be found in Figlio et al. (2015). Much of this literature has focused on the relationship between parental income or education and a child's own level of education. For example, Black et al. (2005) examined the causal impact of parental education on children's education using Norwegian administrative data. More relevant to this paper however, is research by Bhuller et al. (2011) that has considered the impact of education on earnings using Norwegian administrative data on career earnings. They find that conventional estimates of the return to education are downward biased. Carneiro et al. (2013) have also used Norwegian data to investigate the impact of parental income on children's education level and labour market outcomes, finding a strong positive impact from parents' discounted lifetime income on children's outcomes and noting that the timing of changes in family income also makes a difference with evidence of dynamic complementarities. Our paper will contribute to this literature by producing evidence on the variation in graduates' earnings and more specifically the relationship between parental income and graduates' outcomes, using far finer grained measures of the exact nature of graduates' higher education than hitherto.

This paper also builds on previous work using the same dataset in Britton et al. (2015) which focused on describing key features of the distribution of graduates' earnings, and comparing it to that of non graduates. They found that non graduates are twice as likely to have no earnings as graduates, ten years after leaving higher education (30% against 15% respectively for the cohort graduating in 1999 observed in 2011/12). This implies that a degree can provide significant protection from unemployment and non employment. Further, Britton et al. (2015) found that half of non graduate women had earnings below £8k a year at around age 30, while only a quarter of female graduates were earning less than this, and half were earning more than £21k a year. Similar patterns were observed for males. For those with significant earnings (which are defined as above £8k a year) median earnings for male graduates ten years after graduation were £30k, while the equivalent figure for non graduates was £22k. For women with earnings over the £8k threshold, median earnings ten years after graduation were £27k for graduates and £18k for non graduates.

#### 1.4 Structure of the paper

In Section 2 we detail our numerous data sources and in Section 3 our modelling choices. Note that a fuller description of the dataset and its construction can be found in Britton et al. (2015). We

present results showing variation by subject in Section 4 and by institution in Section. We present regression results investigating variation by subject and institution once we control for background characteristics in Section 6 before investigating earnings differences by parental income in Section 7. In Section 8 we conclude and discuss the policy implications of our work.

#### 2 Data and methodology

#### 2.1 Official earnings data

This papers bores into the Britton et al. (2015) anonymised database of official earnings data for English domiciled (at the time they first borrow) borrowers from the Student Loan Company.<sup>2</sup> This covers 10% of all borrowers from 1998 who are in their repayment period (this means they have left HE and a new tax year has subsequently started). Note that throughout we refer to these borrowers as graduates, although it is possible they may not have completed their course. We have HMRC earnings data for each of these individuals who first enrolled in higher education in the academic years 1998-2011 (henceforth known as the 1998-2011 cohorts) and have earnings information for these borrowers from the 2002/03 to 2012/13 tax years. Britton et al. (2015) documented the database and showed results at the country level. Table 1 presents the individual level variables we use in this paper.

Database name	Details	Missing Data
Gender	Female, Male	
First academic year	Date first went to any Higher Education	
	Provider (HEP): 1998 onwards	
Last HEP name	Last HEP attended	Small institutions are grouped together
		and labelled 'OTHER HEP'
Subject group	LEM, OTHER, STEM.	
	LEM denotes Law, Economics & Management	
	STEM denotes science, tech, eng, math	
Subject code	First letter of JACS code	Censored if the n in that year group in
		that subject was less than five.
		Coded as Missing STEM, LEM or OTHER.
Borrowed first year	Given in cash, no interest rate applied	
Region	Government region of address when the	
at application date	SLC application was first made.	
2008/09 earnings	Labour Earnings from PAYE and SA	If there is no tax record
2009/10 earnings	tax forms. All earnings are rescaled to	earnings are coded as 0.
2010/11 earnings	October 2012 prices using the	Legally any income should be
2011/12 earnings	Consumer Price Index (CPI).	reported however small so
2012/13 earnings		zero is the correct (taxable) earning.

Table 1: Individual level variables we use from the Golden sample of the Britton et al. (2015) database of official earnings data for English borrowers.

<sup>&</sup>lt;sup>2</sup>Students in officially recognised higher education learning institutions are eligible for loans. These are defined by the government as either 'recognised' or 'listed', the former can award degrees, the latter can offer courses that lead to a degree from a recognised institution. In our data students studying at both types of institutions will be observed. For example, students achieving their degrees at some Further Education Colleges will be included.

Our information on graduates has some limitations. We have no information on non-English domiciled students, even if they work in the UK, since they do not take out loans from the English part of the SLC. We also cannot identify English domiciled students who chose not to take out a student loan. By the end of the period approximately 85% of English domiciled borrowers had taken out student loans. There is limited empirical evidence to help us understand which individuals do not take out loans, but one might anticipate that students who do not take out loans are likely to be more socioeconomically advantaged, attend higher status institutions and are more likely to go on to be higher earners. We return to this issue later in the paper.

Importantly we only have information on the last HEP the student went to, which means we cannot identify students who swap institutions during their time in HE.<sup>3</sup> For HEPs with fewer than 1,000 loans in total over our 14 years of data the SLC used a blanket rule of recording their institutional name as other HE. other HE is thus a group of smaller HEPs which we will analyse as if they were a single HEP. This impacts around 2% of our data.

Since a principal aim of this paper is to consider the earnings of graduates from lower socioeconomic backgrounds relative to graduates from higher socioeconomic backgrounds, ideally we would have detailed information on parental income. Unfortunately the data do not include a measure of the exact level of parental income. However, the database includes the amount the graduate borrowed in their first year of borrowing. This is useful to us, as the maximum amount the UK Government is willing to loan a student depends upon whether the student is attending a London-based institution and the level of their parents' income, with individuals from lower income households able to borrow more than their more well-off peers.<sup>4</sup>

Based on this, we reverse engineer a measure of parental income. Unfortunately for us the rules are very blunt and there is a lot of noise in the observed amount individuals borrow. However, for each of the 1999-2005 cohorts we observe clear spikes in the distribution of the amount students borrow at the maximum loan levels available to each student from higher income households (we use two of the spikes - the one at the high parental income London maximum and the one at the high parental income non-London maximum). We use this to build a simple binary indicator of the student having higher income parents. This indicator is described in Table 2, with around 20% of individuals appearing at the non-London and London higher income maxima in each cohort.

<sup>&</sup>lt;sup>3</sup>This definitional choice was made because if we had both their first and last HEP then this would be highly disclosive. We opted to have the last HEP, rather than the first HEP, as the former is the institution the student leaves from to join the job market or to carry out further study.

<sup>&</sup>lt;sup>4</sup>This is true for the 1999-2005 cohorts. For subsequent cohorts this rule breaks down, as maintenance grants displaced some loans resulting in a non-monotonic relationship between parental income and loans. The introduction of tuition fee loans from 2006 adds an additional layer of complexity as many poorer students received grants to cover their tuition fees while richer students borrowed to cover their fee loans. We therefore do not use post-2005 cohorts to investigate this measure.

Specifically we define anyone borrowing exactly the amounts given in the table as being from a "Higher income Household". We acknowledge that this measure does not perfectly identify all student from higher income households, for two reasons. First, those from higher income households may borrow less than the maximum available. Second, individuals from low income households may choose to only borrow the higher income person maximum. Each of these factors will bias our estimates of the difference in earnings between the two groups towards zero.

	Min Parental	Loan Am	% higher income identifier						
Cohort	Income	Non-London	London	Overall	Male	Female	$_{\mathrm{LEM}}$	OTHER	STEM
1999	35,000	2,795	3,445	14.6	15.2	14.1	14.0	13.7	16.3
2000	36,000	2,795	3,445	18.9	20.2	17.8	18.9	18.1	20.2
2001	38,500	2,860	3,525	21.4	22.6	20.3	21.5	21.0	21.9
2002	40,000	2,930	3,610	21.8	23.2	20.5	21.7	21.4	22.3
2003	40,000	3,000	3,695	23.8	25.7	22.2	23.5	23.2	25.0
2004	40,950	3,070	3,790	24.8	26.1	23.6	24.1	24.4	25.6

Table 2: Loan amounts used to form the higher income household identifier and share individuals classified as higher income by gender and subject. The parental income column gives the level of income (in nominal prices) above which individuals are eligible for a maximum of the loan amount given in columns 3 and 4 (depending on whether they are attending a London-based HEP).

The earnings data used in this paper are transformed by the Consumer Price Index (CPI) to reflect October 2012 prices. The definition of earnings we use are detailed in Britton et al. (2015) which aimed to record earnings from labour, meaning employment income, profits from partnerships and profits from self-employment are included. Meanwhile, trust income, profits on share transactions, profits from land and property, foreign employment (Britton et al. (2015) remarked they would have liked to have included foreign income but that the calculation involved various delicate deductions that made getting an accurate measure difficult) and savings, UK dividends, pension income, life policy gains, "other" income, bank and building society interest and total income are all excluded, since these variables (except foreign employment) measure non-employment income.

The sample sizes available to us (a 10% sample size of the SLC population database) are given in Table 3, which also shows the gender split and how this varies with the cohort. This shows that for each cohort the larger share of the former student borrowers are female. Importantly, the 1998 cohort is materially smaller than the later ones due to a much lower take up of the Government loan offers. 1998 was the first year of income contingent loans, which replaced the previous offering of mortgage style loans, and it took a year or so for students to adjust to these much less risky loans leading to an increase in the take up rate.

	All England								
Cohort	All	Male	Female						
1998	14,487	6,927	7,560						
1999	22,621	10,590	12,031						
2000	23,506	10,853	$12,\!653$						
2001	23,924	11,025	12,899						
2002	23,891	11,060	12,831						
2003	23,972	11,024	12,948						
2004	23,577	10,767	12,810						
2005	25,103	11,439	13,664						
2006	25,383	11,340	14,043						
2007	25,352	11,292	14,060						
2008	20,847	8,990	11,857						
2009	6,510	3,029	3,481						
2010	2,993	1,334	1,659						
2011	851	360	491						
All	263k	120k	143k						

Table 3: Number of Golden sample (10% sample of loan database) borrowers. Cohort denotes the first year the former borrower received a loan from the SLC. This is a subset of Table 5 in Britton et al. (2015).

Importantly the Student Loan Company data also has the standard Higher Education Statistics Agency "JACS" code for each degree course. To avoid the data being disclosive, the SLC have provided a relatively aggregated version of the JACS code, namely the one digit version. For most students we therefore have the first letter of their JACS code which provides us with 19 separate subject areas. The full list of 19 subjects is at Appendix A.

For some of the smaller institutions this level of subject aggregation might be disclosive when interacted with institution, cohort and gender (the SLC needed at least 5 individuals to be present in each sub-population in order to give us access to the full information). This particularly affects individuals in smaller institutions studying less popular courses. The degree of this missing data will be quantified below. To help us deal with these individuals the SLC coded every individual in the database into three Subject Groups:

- STEM: Subjects in Science, Technology, Engineering and Mathematics
- LEM: Subjects in Law, Economics and Management (Business)
- OTHER: All other subjects.

Economics is coded as JACS code L1, a subset of social studies. SLC broke those individuals out from the rest of social studies and placed them in the subject group LEM with Law (JACS code M) and Management (JACS code N). We designed this LEM grouping of L1, M and N to be subjects which are largely professionally focused. The remaining social studies subjects we placed in the subject group OTHER. Overwhelmingly OTHER is the humanities grouping.

As we have the LEM and OTHER breakdown for every student, when we have a person recorded

as JACS code L we can break them down as L1 (Economics) and LO (other Social Studies). Hence we will analyze these two groups separately in what follows.

Table 4 shows two basic summaries of the subjects students took: first in terms of subject groups and second in terms of the first letter of the JACS code (subject). These are given for two cohorts. In terms of subject groups, LEM is much smaller than the other groups, while OTHER is by far the largest group. OTHER also has the strongest gender bias, with typically just over 60% of the students being female.

Of the individual JACS subject codes, the largest groups are Creative Arts, Education, Biological, Maths and Computer Science and Business. Creative Arts is around 55% female. The subjects, European languages and literature (R) and Other languages and literature (T) are very small and so we have merged these two highly related groups.

	1999 cohort					2002 cohort				
	All	$\mathbf{M}$	$\mathbf{F}$	% Pop	% F	All	$\mathbf{M}$	$\mathbf{F}$	% Pop	% F
	Subject group									
LEM	3,961	1,946	2,015	17.5	50.9	3,760	1,853	1,907	15.7	50.7
Other	11,257	4,243	7,014	49.8	62.3	12,247	4,809	7,438	51.3	60.7
STEM	7,403	4,403	3,000	32.7	40.5	7,884	4,400	3,484	33.0	44.2
					Sub	jects				
JACS A: medicine	424	193	231	1.9	54.5	501	203	298	2.1	59.5
JACS B: allied medicine	1,286	426	860	5.7	66.9	1,141	272	869	4.8	76.2
JACS C: biological	1,125	405	720	4.9	64.0	1,821	721	1,100	7.6	60.4
JACS D: vet & agriculture	227	87	140	1.0	61.7	243	87	156	1.0	64.2
JACS F: physical	867	581	286	3.8	33.0	878	507	371	3.7	42.3
JACS G: math & computing	1,962	1,532	430	8.6	21.9	1,759	1,414	345	7.3	19.6
JACS H&J: engineering and technology	884	775	109	3.9	12.3	952	835	117	4.0	12.3
JACS K: architecture		31					57			
JACS L1: economics	259	205	54	1.1	20.8	221	176	45	0.9	20.4
JACS LO: other social studies	1,692	652	1,040	7.4	61.5	1,764	773	991	7.3	56.2
JACS M: law	1,073	434	639	4.7	59.6	991	340	651	4.1	65.7
JACS N: business	2,428	1,216	1,212	10.7	49.9	2,349	1,242	1,107	9.8	47.1
JACS P: communications	527	213	314	2.3	59.6	912	363	549	3.8	60.2
JACS Q: linguistics & classics	728	201	527	3.2	72.4	779	230	549	3.2	70.5
JACS R&T: languages & lit	257	73	184	1.1	71.6	290	95	195	1.2	67.2
JACS V: history & philosophy	672	316	356	3.0	53.0	825	424	401	3.4	48.6
JACS W: creative arts	2,221	1,010	1,211	9.8	54.5	2,575	1,131	1,444	10.7	56.1
JACS X: education	2,088	532	1,556	9.2	74.5	2,393	629	1,764	10.0	73.7
JACS Missing STEM	578	373	205	2.5	35.5	505	304	201	2.1	39.8
JACS Missing LEM	275	162	113	1.2	41.1	207	114	93	0.9	44.9
JACS Missing Other	2,998	1,175	1,823	13.2	60.8	2,701	1,145	1,556	11.2	57.6

Table 4: Summary of the subject characteristics of borrowers for the 1999 and 2002 cohorts. For every borrower SLC kindly always coded for us the subject group. For small subjects at small HEPs SLC held back the JACS code to protect confidentiality and so there are a group of missing JACS code (although in each individual case we do know, as implied earlier, their subject group). The first 3 columns are counts. The fourth is the percentage of borrowers who are in that group out of the population of all borrowers. The last column for each cohort is the female gender split for those in that group. The last three rows characterise the individuals whose subjects are missing in our database. We can see that they are overwhelmingly in the subject group OTHER.

Table 4 shows the degree of missing subject information, split out by subject group, gender and

cohort. There is very little missing subject information except for in OTHER, reflecting the large number of quite small humanities courses which are taught at smaller institutions.

#### 2.2 Higher Education Statistics Agency data

Here we detail the additional data drawn from the Higher Education Statistics Agency (HESA) student record used to bolster the individual level data from the SLC and HMRC databases. The HESA student record is an administrative record of all students registered at the reporting HE provider (HEP). The HESA student records are provided by individual institutions, following standard definitions established by HESA. HESA then collates an anonymised version of this information in order to provide sector wide reports and sells summaries or subsets of the resulting databases to researchers for secondary analysis. Unless stated we use HESA data from 2002/03. By using a single year of HESA data we are vulnerable to rapid changes in the student profiles of particular degrees and institutions. However, for our analysis of groups of institutions and benchmarking comparisons we need to take a fixed point of comparison.

None of the HESA data can be individually linked to the Britton et al. (2015) database as the HESA records lack any form of identification number through which the SLC and HMRC data could be linked. Instead we use the HESA records to build a profile for the HEP attended and the institution/course the student studied. Table 5 shows the variables which come from the HESA records. For many Subject/HEP combinations the data will be missing as the HEP does not have any courses in the corresponding subject.

We have 11 ethnicity self-identification categories, each of which indicates the proportion of the students taking that Subject/HEP combination self-reporting that particular ethnicity. For each Subject/HEP combination these variables sum to one and so this is also true for each student within our database. To illustrate these features, Table 5 shows the numbers of students in the 2002/03 cohort with each of the ethnicities and the percentage of these which are female. The counts are obtained by summing over each student's ethnicity proportion, while the female count results are only summed over female students.

When we try to control for prior academic achievement in our earnings model we do this using a mean tariff score provided in the HESA data and which in 2002/03 was based on the UCAS "tariff score" at the Subject/HEP level. The tariff score for each individual is a single quantitative summary of the prior performance of students. Each exam passed delivers some grades and the tariff sums them up. A higher tariff indicates more exams passed, typically with higher grades. The Table shows an estimate of the number of students without a tariff score. Most students go into HE having passed various A-levels exams, but some do not, for example those who have taken the international baccalaureate. The database includes the proportion of students without

A-levels. The collection of exams equivalent to A-levels are called Level 3 qualifications. The proportion in the class with Level 3's is also given. Finally, a proportion of missing data are also recorded.

The HESA data include measures of parental occupation, from which a high and low parental occupational class is derived, and the "Participation of Local Areas" (POLAR) classification, which estimates in the student's neighbourhood (roughly, ward level) the proportion of young people who are participating in HE by the age of 19. We use these as measures of deprivation.

The data also have an indicator of the proportion of students in the Subject/HEP who are living at home and whether an individual attended English state schools. As discussed above, the literature has found private school attendance to be an important predictor of both high status university attendance and subsequent earnings. As with all of these HESA data, we do not matched individual level data on this, but instead have the aggregate measure of the proportion in each degree/institution combination who state schools. According to statistics from the Department for Education, in England around a fifth of students in full time education over the age of 16 attended private schools.<sup>5</sup> This suggests that the missing state school indicator in the table includes a large number of privately educated students. Overall, the Parental Occupation and State School indicators are the groups of variables with the most missing data.

<sup>&</sup>lt;sup>5</sup>Department for Education, National tables: SFR16/2015)

					19	999 coho	$\operatorname{rt}$
Topic	Variable name	Type	HEP level	HEP/Course level	Fem	Male	All
Region of HEP							
	Other	Binary	✓		12.1	13.8	12.9
	East Midlands	Binary	✓		8.7	9.2	8.9
	East	Binary	✓		3.4	3.3	3.3
	London	Binary	✓		15.0	13.3	14.2
	North East	Binary	✓		4.9	5.1	5.0
	North West	Binary	✓		12.8	11.9	12.4
	Scotland	Binary	✓		1.2	1.2	1.2
	South East	Binary	✓		12.8	13.0	12.9
	South West	Binary	✓		7.0	7.0	7.0
	West Midlands	Binary	✓		8.3	8.2	8.3
	Yorks & the Humber	Binary	✓		10.5	10.5	10.5
	Wales	Binary	✓		3.3	3.5	3.4
Ethnicity							
	Eth White	Proportion		$\checkmark$	80.0	77.2	78.7
	Eth Indian	Proportion		$\checkmark$	6.1	7.1	6.6
	Eth Bangladeshi	Proportion		$\checkmark$	1.3	1.5	1.4
	Eth Pakistani	Proportion		$\checkmark$	2.8	4.2	3.4
	Eth Chinese	Proportion		$\checkmark$	1.0	1.5	1.2
	Eth OtherAsian	Proportion		$\checkmark$	1.4	1.9	1.6
	Eth BlackCaribbean	Proportion		$\checkmark$	2.4	1.4	1.9
	Eth BlackAfrican	Proportion		$\checkmark$	3.6	3.9	3.7
	Eth BlackOther	Proportion		$\checkmark$	.7	.5	.6
	Eth Other	Proportion		$\checkmark$	3.4	3.2	3.3
	Eth Missing	Proportion		$\checkmark$	19.5	20.7	20.0
Grades	_	_					
	Tariff scores	#			280	273	277
	Tariff score miss	Proportion			19.5	20.7	20.0
	No A-Levels	Proportion			16.5	15.3	15.9
	No Level 3	Proportion			14.3	12.6	13.5
	Hiqual miss	Proportion			19.5	20.7	20.0
Parental occupation							
	Class high	Proportion		$\checkmark$	42.9	42.9	42.9
	Class low	Proportion		$\checkmark$	25.5	28.0	26.7
	Class miss	Proportion		$\checkmark$	19.5	20.7	20.0
Deprivation index		-					
=	Polar	Proportion		$\checkmark$	10.4	9.7	10.1
	Polar miss	Proportion		$\checkmark$	19.5	20.7	20.0
Living at home		•			23.6	23.4	23.5
School type							
v <u>*</u>	State school	Proportion		$\checkmark$	91.0	88.3	89.7
	State miss	Proportion		$\checkmark$	19.5	20.7	20.0

Table 5: HESA based information for each individual. The Table shows the variables, the character of the data and if the data are taken at the Higher Education Provider (HEP) level or the HEP/course level. The notation "hiqual" means either A-levels or equivalent. Proportions are recorded on the interval [0,1]. The low share of individuals at state school is likely because many individuals not from state schools are classified as "unknown or not applicable". Although they are indistinguishable from the genuinely missing, this suggests the "State miss" variable is informative. NOTE: this table is currently incomplete as we are awaiting additional data.

#### 2.3 Data limitations

As we have already indicated, only those who borrow from the SLC are included in the sample. This excludes those who choose not to take out a loan. Further, only higher education institutions whose students are eligible to receive a loan from the SLC are included. This will exclude students

doing some tertiary level courses in Further Education colleges, for example, who do not qualify for loans.

In terms of data quality, our earnings data are of unparalleled quality and size. We are also confident that the institution of study variable will have a high degree of accuracy. For very small specialist institutions, institutions will be coded "otherHE". We have also obtained permission from a subset of universities to be named in our analysis. We invited all members of the Russell Group to be named in the analysis as an initial feasible first step, though later we would want to invite all universities to participate. Thus far the following institutions have kindly agreed to this: Queen's Belfast, Bristol, Cambridge, Cardiff, Durham, Edinburgh, Exeter, Glasgow, Imperial, King's, Liverpool, LSE, Manchester, Oxford, Newcastle, Nottingham, Southampton, York and Warwick. Unfortunately Glasgow and Queen's Belfast have quite small sample sizes of English domiciled borrowers in our 10% sample and so we do not name them in this version of the paper as the results would be unreliable - we will do so in an updated draft which we hope will draw on the 100% sample (of English domiciled students) from the administrative data. We also note that the data for Cardiff and Edinburgh (in particular) are not necessarily representative of the bulk of their graduates since we only have data on England domiciled students.

	Average		1999	cohort		2002 cohort				
	Tariff	All	M	F	% F	All	M	$\mathbf{F}$	% F	
Cambridge	501	254	120	134	52.8	225	101	124	55.1	
LSE	446	39								
Newcastle	375	205	94	111	54.1	197	94	103	52.3	
Nottingham	427	290	137	153	52.8	364	188	176	48.4	
Oxford	502	247	128	119	48.2	198	109	89	44.9	
Southampton	357	272	136	136	50.0	238	106	132	55.5	
Warwick	419	179	88	91	50.8	234	111	123	52.6	
Exeter	342	180	93	87	48.3	213	112	101	47.4	
York	417	100	49	51	51.0	128	68	60	46.9	
Liverpool	328	218	99	119	54.6	252	121	131	52.0	
Durham	425	213	99	114	53.5	266	111	155	58.3	
Bristol	420	205	100	105	51.2	228	94	134	58.8	
Cardiff	367	146	60	86	58.9	190	90	100	52.6	
Edinburgh	390	102	46	56	54.9	92				
King's College	370	165	86	79	47.9	200	86	114	57.0	
Manchester	370	418	217	201	48.1	436	210	226	51.8	
Imperial	474	94	61	33	35.1	106				

Table 6: Summary of case studies. Figures have been suppressed in cases of sample sizes of fewer than 30 individuals (and in cases where a small sample size could be inferred from other information in the row).

Table 6 shows basic descriptive statistics for each institution that we name, including the mean tariff score for each institution and student counts. The variation in the prior achievement levels of students enrolled in each institution is obviously critical and we return to this issue below. Note also that we use multiple years of tax data (5) and cohorts (7), increasing sample sizes typically by a factor of around 35.

There are of course many other higher education institutions and without explicit permission we cannot name them in the analysis<sup>6</sup>. To overcome this problem and to make the analysis and interpretation of our research more tractable we use typologies of universities. The HESA data does include a typology of HEPs, namely self-declared mission groups, such as Millionplus or the Russell Group. For the purposes of analysing earnings variation this may not necessarily be the optimal way of classifying institutions and in any case the membership of these mission groups changes over time. We therefore do not use these groupings, and instead where appropriate group HEPs according to the mean UCAS tariff score of their student intake, dividing the population of institutions into deciles on the basis of their undergraduate population mean tariff score on entry. For some analyses we also split the upper decile into two groups in recognition that there is a distinct group of institutions at the very top of the distribution which may be of particular interest. Table 7 shows the groupings, the numbers of males and females in each group, the percentage of each cohort in each group and the proportion in each group that is female. This is relatively stable across groups except for G10high where it dips below 50%. Note there is a large group of institutions that are placed in group zero because they have missing tariff information.

HEP	#	Av.		1999 cohort					2	002 coh	ort	
Group	HEPs	tariff	All	Male	Fem	%Pop	%Fem	All	Male	$\operatorname{Fem}$	%Pop	%Fem
G0	62		4,528	2,188	2,340	18.5	51.7	4,731	2,270	2,461	18.2	52.0
G1	12	175.3	2,369	1,076	1,293	9.7	54.6	2,107	935	1,172	8.1	55.6
G2	11	201.2	1,674	730	944	6.8	56.4	1,652	779	873	6.4	52.8
G3	11	216.7	1,739	771	968	7.1	55.7	1,672	687	985	6.4	58.9
G4	12	232.3	1,916	884	1,032	7.8	53.9	2,127	977	1,150	8.2	54.1
G5	11	250.5	2,172	1,050	1,122	8.9	51.7	2,432	1,125	1,307	9.4	53.7
G6	11	267.6	1,343	599	744	5.5	55.4	1,378	626	752	5.3	54.6
G7	12	295.6	1,270	599	671	5.2	52.8	1,549	752	797	6.0	51.5
G8	11	342.0	1,632	788	844	6.7	51.7	1,738	803	935	6.7	53.8
G9	11	370.9	2,094	977	1,117	8.5	53.3	2,414	1,067	1,347	9.3	55.8
G10	11	440.0	1,884	930	954	7.7	50.6	2,091	1,041	1,050	8.0	50.2
G10low	6	417.9	1,240	591	649	5.1	52.3	1,525	735	790	5.9	51.8
G10high	5	492.5	644	339	305	2.6	47.4	566	306	260	2.2	45.9

Table 7: University groupings for the 1999 and 2002 cohorts. Group G0 is the group of institutions for which there is no tariff score available.

In the UK earnings vary substantially by region and a limitation of our data is that we do not have the area in which the graduate is currently located (in any case graduates' locations may be considered endogenous with graduates who have more human capital being more able to relocate to higher earning regions). However, the database contains the Government region of the borrower on application, which is strongly correlated with the region in which the HEP lies and hence the graduates' current location, since we know that a high proportion of graduates remain in their

 $<sup>^6</sup>$ Under UK Parliamentary Statute, HMRC treats institutions as individuals whose privacy is protected unless they explicitly wave their confidentiality.

region of study. Inclusion of the region of the HEP in the model will therefore take some account of the fact that institutions located in regions with weaker labour markets may, through no fault of their own, have graduates who earn less. This region variable is quite coarse, dividing the UK into 12 areas (see Table 5 above). For each HEP and subject combination we can also include the percentage of students who live at home whilst studying, which will also provide an imperfect indicator of whether students taking that degree are likely to be mobile to develop their career.

							%Own
	All	$\mathbf{M}$	$\mathbf{F}$	% Pop	% F	%G10	Region
East Midlands	1,515	686	829	7.5	54.7	11.0	46.3
East	2,067	954	1,113	10.2	53.8	12.1	19.8
London	3,412	1,602	1,810	16.8	53.0	8.1	59.8
North East	922	416	506	4.5	54.9	8.7	73.5
North West	2,976	1,381	1,595	14.7	53.6	6.5	62.0
South East	3,462	1,635	1,827	17.1	52.8	11.8	38.3
South West	1,980	927	1,053	9.8	53.2	8.3	41.5
West Midlands	2,194	1,007	1,187	10.8	54.1	7.9	49.3
Yorks & the Humber	1,746	797	949	8.6	54.4	9.8	55.6

Table 8: 1999 cohort. %G10 is the percentage of students from that region that go to top 10% institutions ranked by Tariff score. %Own region is the percentage which studied in the same region as the region from which they first applied for a loan (i.e. typically their home before they went to study).

To understand the significance of region, Table 8 shows the regions in which the 1999 cohort of students lived when they applied for a student loan. The first three columns of data show the counts for students by region of origin. The next column shows the proportion of the sample that is originally from each region, with London, the South East and the North West having the highest percentages of the total sample of students. The next column shows the percentage of students from each region who are female, showing little variation by region. The penultimate column indicates the proportion of students from each region who are enrolled in our top decile of institutions (Group G10). Participation in these top institutions is somewhat higher in the South East, the East and the East Midlands regions. What is particularly striking is the low participation in these top institutions amongst students from the North West, and to a lesser extent, the West Midlands. Some of these differences reflect the geographic distribution of these top institutions, and is consistent with previous evidence suggesting proximity to a university influences participation. The final column gives the percentage of students from each region who attended a university in their own region. This too varies significantly across regions: nearly three quarters of those from the North East attended an institution in that region whilst only one fifth of those from the Eastern region attended a university there. This too partially reflects the geographic distribution of institutions though it will also reflect decisions to stay at home whilst studying at university, with students from poorer backgrounds more likely to study close to home.<sup>7</sup>

 $<sup>^7\</sup>mathrm{Authors}$ own calculations using linked National Pupil Database (NPD) - HESA data.

A final limitation is that we would ideally like to model graduate' earnings throughout their life until retirement. However, given that administrative data on student loans does not go that far back in time, for most graduates our database is likely to cover the first ten years or so of their careers.

#### 3 Modelling

We will model how earnings vary by individuals' characteristics, namely age, gender, indicators of socioeconomic background and of course subject of degree and institution of study. Our focus will be on quantiles.

The model employed to estimate the earnings distribution will be calibrated from data on the 1999-2005 cohorts and for the five tax years which run 2008/09-2012/13. Typically we report results for the 2012/13 tax year for the 1999 cohort.

The model structure we use has

$$\tau = \Pr(Y_{i,t} < q_{it}(\tau)|\mathcal{Z}_i) \tag{1}$$

where  $Y_{i,t}$  is the earnings for person i at time t,  $\mathcal{Z}_i$  are conditioning variables known about person i from the SLC database at time of first application for a loan (e.g. cohort, gender, year).

Here  $\tau$  is the quantile level and  $q_{it}(\tau)$  is the model based quantile, where

$$q_{i,t}(\tau) = \beta_0(\tau) + \beta_1(\tau)Fem_i + \beta_2(\tau)Cohort + \beta_3(\tau)Cohort^2$$
(2)

$$+\beta_4(\tau)Cohort_i \times Fem_i + \beta_5(\tau)Cohort_i^2 \times Fem_i + \gamma(\tau)'t$$
(3)

and Fem is a female dummy, and Cohort is set equal to 0 for individuals who first went to university in 1999, increasing by 1 with each year. t has a set of year dummies  $\gamma(\tau)$ .

This model is estimated using a quantile regression at the  $100\tau \in 10$ , 20, 30, 40, 50, 60, 70, 80, 90, 95, percentiles, and the plots show the predicted averages at each percentile.

A central modelling problem will be the extent to which the earnings profiles that we construct can be used as estimates of the relative economic value of different higher education options, specifically the wage benefit from studying at a particular institution or taking a particular subject as compared to another HE option. As already discussed, we will address this problem by using additional aggregate data, particularly from the Higher Education Statistics Agency (HESA), to relate the earnings differences that we observe to differences in the prior achievement levels of the student intake as measured by the average HESA tariff point score for the particular degree subject-institution combination. This is an approach that is similar to many typical value-added models that are used to measure school effectiveness, whereby controls for prior achievement are added

to a model which is attempting to determine the importance of schools in explaining achievement of pupils, where it is important to account for differences in pupil intake (for example, Ladd and Walsh (2002)). Using this model, we can ascertain whether some institutions appear to produce graduates with earnings that are exceptionally high or low in comparison with other institutions with similar student intakes. Whilst this approach will not enable us to completely overcome the problem of ability bias, it does mean we can be more sure that we are comparing the earnings of graduates with similar levels of prior educational achievement. We will also use this approach to allow for other factors that might influence graduate earnings, such as the demographic profile of the student intake, the socioeconomic profile of the students and the location of the institution since location is a key determinant of earnings.

#### 4 Variation by subject

#### 4.1 Three case studies for subject

We start by illustrating the advantages of our data and our method of analysis by considering some individual subject case studies.

We initially focus entirely on three example subjects (each of which is large when aggregated over the genders) for the 1999 cohort. Creative Arts (there share of which has grown strongly in English HEPs) and Business Studies each have about 10% of students. Mathematics and Computer Sciences has around 8%. We will give results for each of these subjects, and contrast them with results for all university students and with those who did not go to HE (whose aggregate results are taken from Britton et al. (2015)). Throughout we report estimated percentiles of the distribution. On the right hand side of the picture, beyond the red line, we also display the mean of the subpopulation. The cross is the mean for the non-university people. The pictures display earnings on a log scale, which will magnify differences in those with low earnings but allow us to see proportional differences in earnings levels.

Figure 1 shows annual earnings for female graduates who studied selected degree subjects. The results are for the 1999 cohort of female graduates in the tax year 2012/13, and these graduates' earnings are presented at various percentiles in the earnings distribution. These individuals have been in the labour market for around a decade. Note the **figures include individuals with no reported earnings** and so are not comparable to many data sources which often present estimates of earnings for those graduates who are in full time employment only. We have found that it is far simpler and more transparent to report the entire distribution as labour market participation rates vary dramatically between different sub-populations we study and using percentiles shows this up in a clear all encompassing way. Figure 2 shows similar information for male graduates.

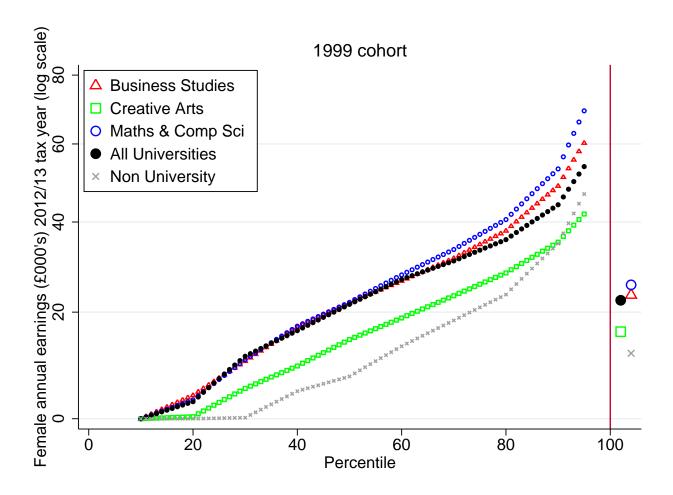


Figure 1: Quantiles of female earnings for the 1999 cohort from three subjects compared to the quantiles for all female graduates. No control variables are used for intake. Discrete points are taken from the distribution, at the 10th, 20th... 90th, 95th quantiles, with linear interpolation in between. This may give the impression of understating the share with zero earnings, for example. Scatter points to the right of each Figure show the corresponding mean for each case (the horizontal positioning of the dots is entirely random, this added jitter makes the dots easier to read). Note earnings are displayed on a log scale.

A significant observation from these pictures is that graduates have a much smaller share of people with very low earnings. Around 20-30% of non-graduates have no earnings, while the equivalent figure for graduates is around 15%. This holds for each subject and gender.

Though this figure may seem high, there are a number of reasons for an individual to have no earnings in these data in addition to simply being out of work. These include undertaking further study, tax avoidance, moving abroad and death. The share of low earners is also impacted by tax allowances, the self-employed claiming offsetting costs against income and other forms of income that we do not include (since our focus is on earned income). Our findings are supported by Student Loan Company official statistics which show 9% of those still in repayment from the

1999 cohort with no employment.<sup>8</sup> A further 1.4% are recorded as having had their loans written off due to death, disability or bankruptcy, many of which are likely to still have zero earnings in 2011/12, but would not be incorporated in the 9% figure. Meanwhile, some of the 37% who have cleared their debt by this point will have zero earnings, while we estimate that approximately 2% will have zero earnings in our data due to moving abroad. We therefore believe our results to be an accurate reflection of UK reported earnings. This latter claim is further supported by our previous work which shows the earnings distribution of graduates in these data is reassuringly similar to that observed in the Labour Force Survey Britton et al. (2015).

Overall both panels shows some variation in graduates' earnings across the different subjects and in particular what stands out is the relatively low earnings of graduates in the Creative Arts across the distribution. More than half of their graduates have earnings below £20k for both genders. Male Creative Arts graduates particularly struggle delivering average earnings which are roughly the same as non-graduates (later we will confirm that of all subjects Creative Arts is the one which delivers, for both men and women, the lowest results for earnings by some margin). Mathematics and Computer Sciences and Business are roughly the same, with both doing a little better than the results for graduates as a whole, who in turn earn far more than non-graduates (grey crosses). The premium for studying these two subjects over other graduates seems higher for women than men.

A further important issue which is prompted by these Figures is that within each subject there is considerable variation in graduates' earnings. A key question is the extent to which variation is systematically related to institution attended, the grades needed to get into that subject (e.g. Mathematics and computing is typically taught to students with higher tariff scores), etc. This issue is addressed later in the paper.

<sup>&</sup>lt;sup>8</sup>These data are not perfectly equivalent, however, as SLC only consider those who have been observed in the UK tax system and they include EU borrowers.

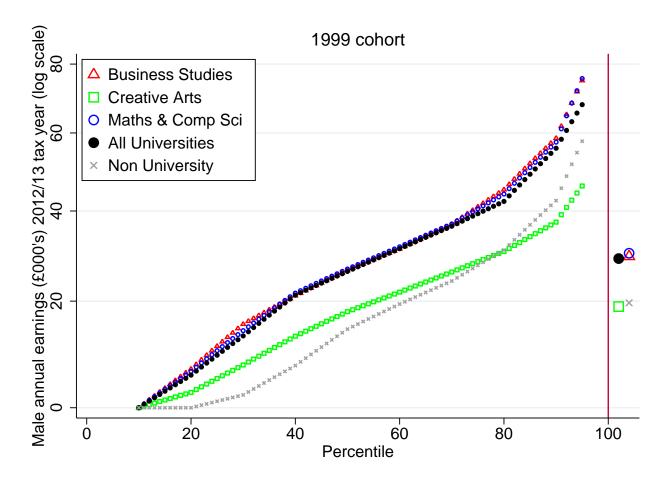


Figure 2: Quantiles of male earnings for the 1999 cohort for three subjects compared to the quantiles for all male graduates. No control variables are used for intake. Discrete points are taken from the distribution, at the 10th, 20th... 90th, 95th quantiles, with linear interpolation in between. This may give the impression of understating the share with zero earnings, for example. Scatter points to the right of each Figure show the corresponding mean for each case (the horizontal positioning of the dots is entirely random, this added jitter makes the dots easier to read). Note earnings are displayed on a log scale.

#### 4.2 All cases for subject

We now move to taking a wider view, looking at all the different subjects recorded in the data. We report the 20th, 50th and 90th percentile of graduates' earnings for each subject, split by gender. Figure 3 shows the results for females. We rank subjects by median earnings, and graph from the lowest median earnings to the highest. Figure 4 shows the results for men. For women it is apparent that for lower earners (at the 20th percentile) there is little variation in earnings by subject. By contrast at the 90th percentile the variation in earnings across subject is more evident with graduates of some subjects, such as economics, medicine, law and languages, going on to achieve significantly higher earnings than other subjects. Patterns are quite similar for males, though the variation in earnings is greater. For example, males from this cohort whose earnings were at the

90th percentile and who studied economics earned in excess of £120k, whilst those who studied Creative Arts earned less than £40k at the 90th percentile. However, caution is needed when interpreting these results since they take no account of the student intake into different subjects and the fact that some subjects attract students with much higher levels of prior achievement.

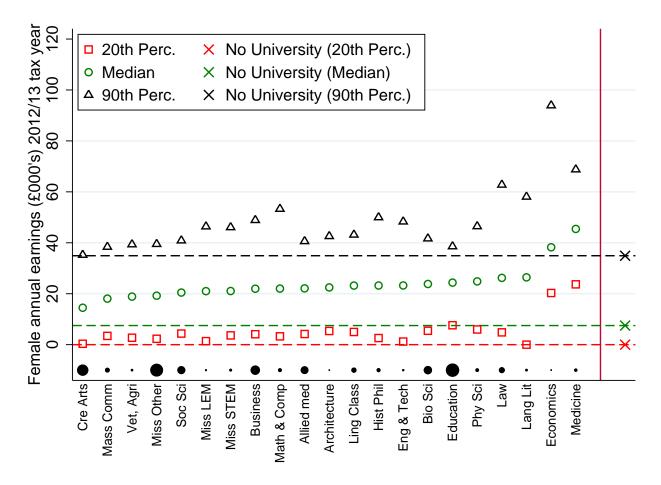


Figure 3: 20th, 50th, 90th percentile earnings for 1999 cohort female graduates in 2012/13 by subject. The area of the solid blob indicates subject size. The crosses to the right of each Figure shows earnings at the 20th, 50th and 90th percentiles across all subjects.

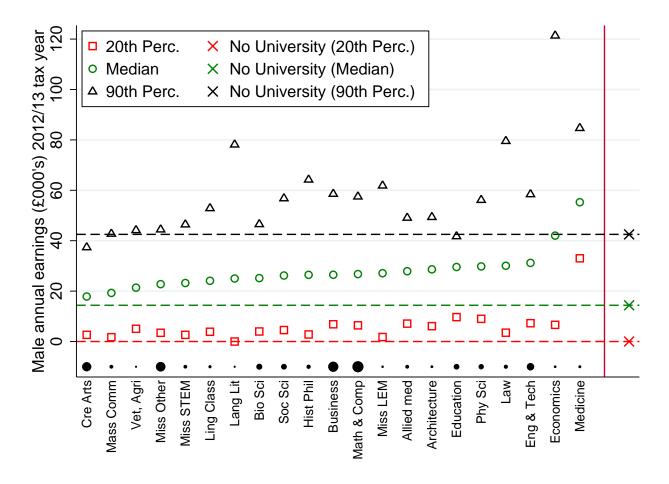
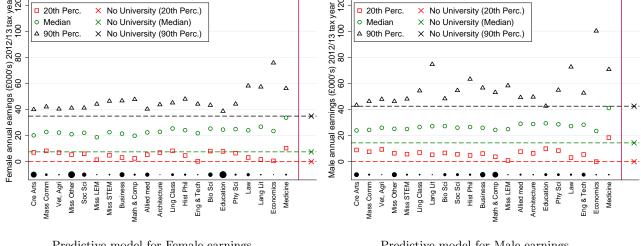


Figure 4: 20th, 50th, 90th percentile earnings for 1999 cohort male graduates in 2012/13 by subject. The area of the solid blob indicates subject size. The crosses to the right of each Figure shows earnings at the 20th, 50th and 90th percentiles across all subjects.

We now consider some conditional estimates that do take account of differences in student characteristics across subjects. To account for prior achievement of students the left hand side of Figure 5 shows the earnings of females by subject but this time conditional on other factors that influence earnings, including age, region, parental income and the full set of HESA characteristics (the method we use for this is described in more detail below). The latter are at subject-institution level and include tariff score of intake. The right hand side of Figure 5 contains the results for men. Once some account is taken of student and course characteristics, the variation in graduates' earnings are reduced somewhat. Nonetheless the main findings still hold. There is little variation by subject at the 20th percentile for males or females. At the 90th percentile subject matters more for both genders and in particular graduates of medicine, law, economics and languages continue to go on to achieve much higher earnings. However, it should be noted that we believe the model we use for our conditioning is more accurate at the median than in the tails, meaning the 90th and 20th percentiles here should be treated with caution (the same does not apply for the unconditional

plots above).



Predictive model for Female earnings

Predictive model for Male earnings

Figure 5: Model based predicted earnings at the 20th, 50th and 90th percentiles for the 1999 cohort in 2012/13 by subject. Left is female, right is male. The predicted model includes controls for university characteristics at the subject-institution level, age, parental income and region of the student, in addition to the control for year of observation. The solid blob indicates the number of students taking that particular subject. The crosses to the right of each Figure shows earnings at the 20th, 50th and 90th percentiles across all subjects.

#### 5 Variation by institution

#### 5.1Some case studies of institutions

To start we again focus on three examples, a subset of the English universities that have kindly given us permission to name them in our study. Our choices are University of Cambridge, Southampton University and Warwick University.

In Figure 6 we show for each institution, annual female earnings for the 1999 cohort of graduates in the tax year 2012/13 at each percentile in the earnings distribution. The earnings for graduates from each institution are shown separately, along with earnings for graduates from all institutions included for comparison (black dots), together with the earnings of non-graduates using the grey crosses. There is considerably more variation across these institutions than we saw across the three subjects in the previous section (note carefully the different scales).

The scatter points to the right of each figure show the mean annual earnings of graduates from each institution and for all institutions. Overall it shows, of the three HEPs used here, graduates from the University of Cambridge have the highest earnings for the upper part of the earnings distribution, with more bunching across institutions at the 50 percentile level. There is much more variation at the higher quantiles. The gaps between the universities seem more pronounced for men than for women (recall the figures are drawn on the log scale), an effect which we will see holds up for a wider set of HEPs.

To take an important snapshot, at the median earnings within each subgroup for females: non-university earnings are £7.5k, all university earnings are £21.6k, while the universities highlighted here scatter between £24.8k-£32.5k, including individuals with zero earnings in the calculations. Hence, roughly, female graduate earnings are three times higher than non-university earnings, and this group of institutions boosts earnings over all universities by another factor of 1.1-1.5.

At the 90% percentile the corresponding numbers are £34.9k, £44.2k and £50.8k-£71.5k. Hence, the multiples are 1.3 and 1.1-1.6.

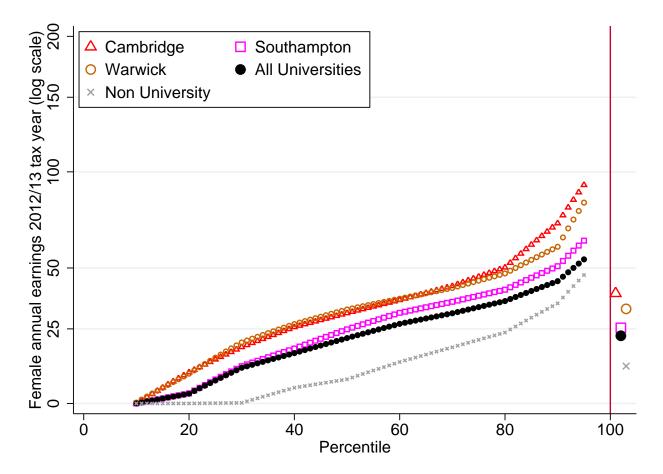


Figure 6: Quantiles of female earnings for the 1999 cohort from three of our case study universities compared to the distribution for all female graduates. Discrete points are taken from the distribution, at the 10th, 20th... 90th, 95th quantiles, with linear interpolation in between. This may give the impression of understating the share with zero earnings, for example. Points to the right of each figure show the mean for each case (the horizontal positioning of the dots is entirely random, this added jitter makes them easier to read).

Figure 7 gives the corresponding results for men. Again the University of Cambridge has higher earnings over the upper part of the distribution. If we focus on the median earnings within each subgroup for males: non-university earnings are £14.4k, all university earnings are £26.3k, while the universities highlighted here scatter between £29.5k-£38.7k. Hence, roughly, male graduate earnings are roughly twice as high as non-university earnings, and this group boosts over all uni-

versities by another (multiplicative) factor of 1.1-1.5 (which very similar to the boost for women).

At the 90% percentile the corresponding numbers are £42.5k, £55.9k and £71.6k-£121.4k. Hence, the multiples are 1.3 and 1.7-2.9, much greater than the corresponding multipliers for women.

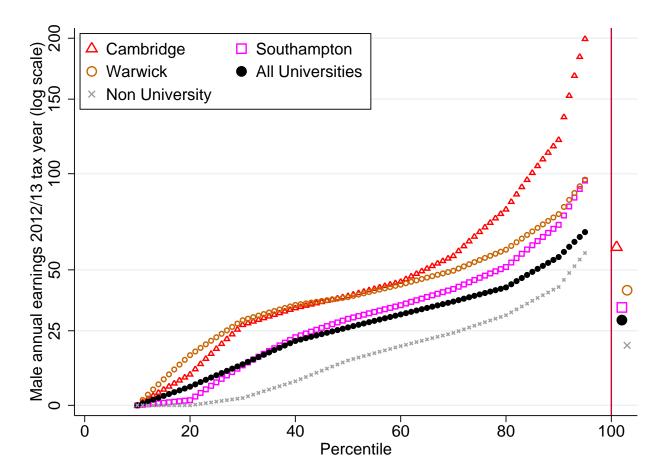


Figure 7: Quantiles of male earnings for the 1999 cohort from three of our case study universities compared to the distribution for all male graduates. Discrete points are taken from the distribution, at the 10th, 20th... 90th, 95th quantiles, with linear interpolation in between. This may give the impression of understating the share with zero earnings, for example. Points to the right of each figure show the mean for each case (the horizontal positioning of the dots is entirely random, this added jitter makes them easier to read).

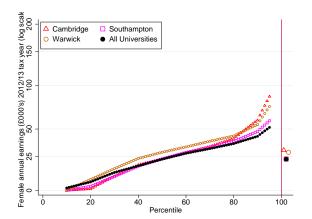
Each of the three named institutions have graduates with higher earnings than the average for graduates from all institutions, at least from the 40th percentile upwards. This observation is not surprising, as we have named institutions that have a more academically selective intake and therefore are likely to admit students who would have better earnings prospects than average, irrespective of their higher education institution. Further, within the sample of institutions there is variation in student intake, the balance of subjects offered and region in which the HEP is located.

All these factors can dramatically influence the unconditional earnings distributions that we show in Figures 6 and 7 and might make direct comparisons of institutions using the level of earnings of students misleading if we are interested in the value added by each institution. For example, graduates from Cambridge have the highest earnings for the upper parts of the earnings distribution. This is likely to be partially explained by its economically advantageous location and highly selective student intake.

To start to address this issue we build a predictive model for the earnings distribution at each institution by gender. This is done in two-step process. First, we "correct" the earnings for each institution using a pooled regression of earnings on a set of demeaned characteristics including subject studied, region applied from, whether or not the individual is from a higher income household and the HESA controls. Our corrected earnings are then the residuals from this regression. Second, we re-estimate the model described in section 3 with  $Y_{i,t}$  replaced with corrected earnings.

Figure 8 shows these corrected earnings for each of our three case study universities. The figure shows that much of the differences between institutions is expected when we account for the differences in these background variables that influence earnings. This is perhaps most obvious when looking at the scatter plots of means which indicate the compression in mean annual earnings differences across institutions in the conditional plots. In other words, mean differences in earnings across most institutions are not sizeable once we take account of the fact that different types of student sort into different institutions.

Of course, this exercise is limited. In particular, the tariff scores are pretty coarse measures of pre-HE performance, particularly at the higher level of achievement. Admissions officers at HEPs have a broader set of quantitative and qualitative information with which to select students. This means our conditional approach is likely to be only an approximate way to calibrate expected performance of HEPs. Further, we again note that we believe the model we use for our conditioning is more accurate towards the middle of the distribution than in the tails.



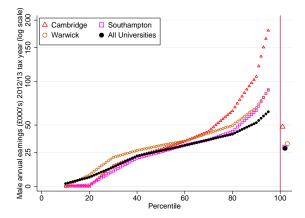


Figure 8: Model based predicted earnings for the 1999 cohort from three of our case study universities compared to the distribution for all graduates, by gender. The predicted model uses an additive model based on the variables: the mean UCAS tariff of the institution, parental income, subject dummy variables and region of institution. Discrete points are taken from the distribution, at the 10th, 20th... 90th, 95th quantiles, with linear interpolation in between. This may give the impression of understating the share with zero earnings, for example. Points to the right of each figure show the mean for each case (the horizontal positioning of the dots is entirely random, this added jitter makes them easier to read).

#### 5.2 All institutions

Figure 9 shows the 20th, 50th and 90th percentile earnings for the female 1999 cohort observed in 2012/13 displayed for each institution for which we have sufficient data to report estimates. Figure 11 shows the same information for male graduates. Universities are ranked left to right on their graduates' 2012/13 median earnings (hence the green line of dots looks smooth). This shows unconditional earnings by institution at the different parts of the earnings distribution.

Our named institutions are included for illustrative purposes.<sup>10</sup> They are displayed with a fuller colour and are numbered, while we have faded the colours from the other HEPs. Each of these named universities have results which appear in the top third of our median income ranking. They are magnified in Figures 10 and 12 which focuses on this top third to make it easier to read. One notable observation is the very strong performance of some northern universities, Liverpool, Newcastle and York, which have graduates that achieve highly competitive earnings even though their local labour markets have lower earnings than we see in the southern part of England. Also note the strong levels of earnings of the London based HEPs: Imperial, LSE and Kings. LSE additionally benefits from focusing also on high paying subjects, Economics and Law, as well as having very high admissions requirements. Taken together, both female and male LSE graduates

<sup>&</sup>lt;sup>9</sup>This should not be considered as comparable to the numerous rankings of universities that are currently available in the UK, and is simply used to make the Figures easier to read. There is no attempt here to allow for subject of student composition, for example.

<sup>&</sup>lt;sup>10</sup>We reiterate here that our estimates are for those domiciled in England upon application. This is a select sample, therefore, in particular for Cardiff and Edinburgh universities, as the majority of their students were domiciled upon application in Wales and Scotland respectively.

are some of the very highest earnings graduates in the country.

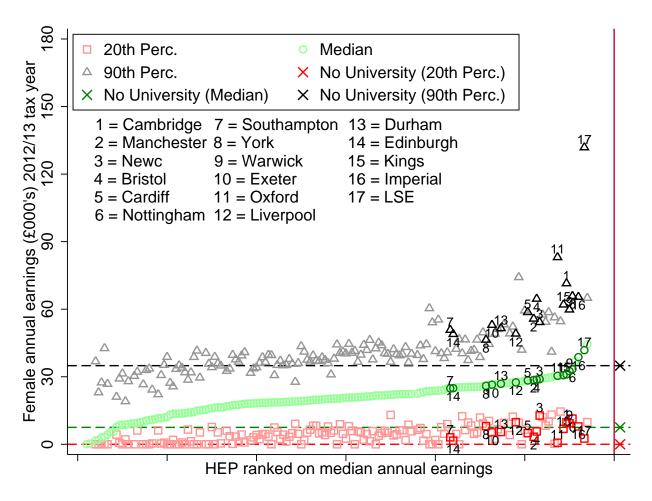


Figure 9: Unconditional female 20th, 50th and 90th percentile earnings for the 1999 cohort in 2012/13 for HEPs ranked on their graduates' 2012/13 median earnings. There are 166 different institutions included, and one "other" institution which include several hundred institutions that issue only a handful of loans. Note: The log scale is not used here. Zeros are included.

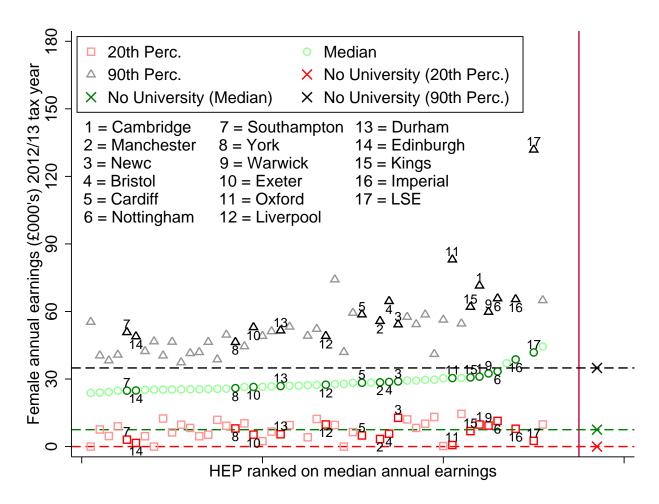


Figure 10: Top third of Figure 9. Unconditional female 20th, 50th and 90th percentile earnings for the 1999 cohort in 2012/13 for HEPs ranked on their graduates' 2012/13 median earnings. This is a repeat of Figure 9, zoomed in on the top one third of HEPs. Note: The log scale is not used here. Zeros are included.

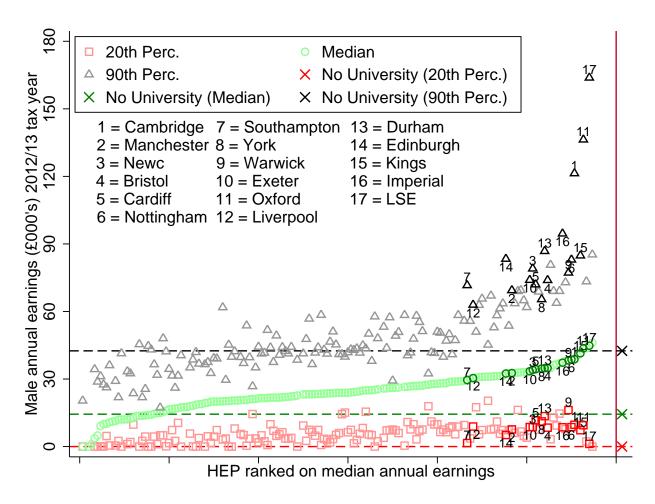


Figure 11: Unconditional male 20th, 50th and 90th percentile earnings for the 1999 cohort in 2012/13 for HEPs ranked on their graduates' 2012/13 median earnings. There are 168 different institutions included, and one "other" institution which include several hundred institutions that issue only a handful of loans. Note: The log scale is not used here. Zeros are included.

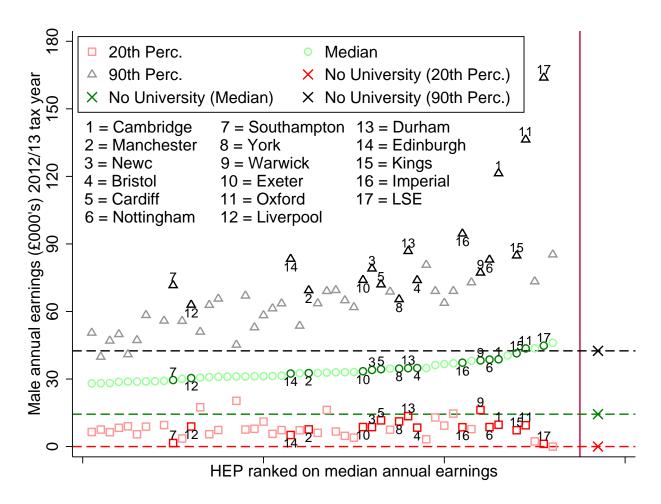


Figure 12: Top third of Figure 11. Unconditional male 20th, 50th and 90th percentile earnings for the 1999 cohort in 2012/13 for universities ranked on their graduates' 2012/13 median earnings. This is a repeat of Figure 11, zoomed in on the top one third of HEPs. Note: The log scale is not used here. Zeros are included.

Figures 10 and 12 repeat the previous figures but focus solely on the top third of institutions to make reading the figures somewhat easier.

To provide some context for the display of graduates' earnings, the average for non-graduates at the 20th, 50th and 90th percentiles is indicated by a cross and associated dashed line. What is perhaps most interesting is the sheer quantity of variation in graduates' earnings within an institution. The earnings of male graduates from the LSE for example are around £163k at the 90th percentile but at the median are nearer £45k while at the bottom 20th percentile the figure is close to zero. The figures also illustrate the large gender gap prevalent for many institutions, particularly at the top end of the earnings distribution.

Another striking feature of the data is the very low earnings of graduates from most institutions at the 20th percentile of the distribution. We noted already that these data include graduates with zero and low earnings, including those working part time or indeed not working at all (whether by choice, e.g. taking more training, or due to unemployment). A significant proportion of the

graduate labour market is therefore low earning (noting that graduates who are abroad appear as having zero earnings in our data), irrespective of institution, but in these data we are unable to say why that might be. However, to put this in context, this low earning share is lower than we see in the non-university population.

Note that we observe a similar pattern for males and females and hence this feature of the data cannot be entirely explained by women choosing to stay out of the labour market in larger numbers. Given previous evidence that lower income students are less likely to access Russell Group institutions which appear to be higher ranking in terms of their graduates' earnings, a major issue for those concerned with improving social mobility is the extent to which students from lower income families are disproportionately likely to be found in these groups of much lower earning graduates.

Another striking feature is the substantial earnings of graduates at the 90th percentile across a large number of institutions. Graduates from all our named institutions earn £60k or above at the 90th percentile of the distribution and even at very low ranked institutions graduates at the 90th percentile earn £30k or more, even including graduates with zero earnings. Further these data illustrate the earnings advantage of graduates as compared to non-graduates. Almost all institutions have graduates with earnings above the 20th percentile of the non-graduate earnings distribution, and most institutions have graduates with earnings above the non-graduate median.

At the other end of the spectrum, there were some institutions (23 for men and 9 for women) where the median graduate earnings were less than those of the median non-graduate ten years on. It is important to put this in some context though. Many English higher education institutions draw a significant majority of their students from people living in their own region. Given regional differences in average wages, some very locally focused institutions may struggle to produce graduates whose wages outpace England-wide earnings, which include those living in London etc. To illustrate regional differences, employment rates in the period under consideration varied between 66% in the North East and 75% in the East of England, and data from the Annual Survey of Hours and Earnings suggests that average full-time earnings for males were approximately 48% higher in London than in Northern Ireland, and around 34% higher for females. Regional differences are therefore important and we take them into account in our analysis of graduates' earnings. However, we cannot construct a more natural benchmark for these locally focused institutions, such as an estimate of the quantiles of the earnings of non-university people in their region, because the data we received from HMRC on non-graduates do not have that regional indicator and so we are unable to carry out that comparison.

As discussed earlier we are conscious that the earnings of graduates from different institutions

will vary because of the student intake or subject mix at that particular institution. Figure 13 provides the predictive estimates of differences in graduates' earnings by institution and suggests significant compression of the differences by institution once account is taken of other factors that influence earnings, such as student intake. Certainly institutional variation at the 20th and 50th percentiles is much reduced. There continues to be outlier institutions with very high earnings at the 90th percentile, although we again note that we believe the model we use for our conditioning is more accurate at the median than in the tails, meaning the 90th and 20th percentiles here should be treated with caution (the same does not apply for the unconditional plots above). It also shows the rankings of the named institutions remain roughly similar even after controlling for differences in other variables that influence earnings.

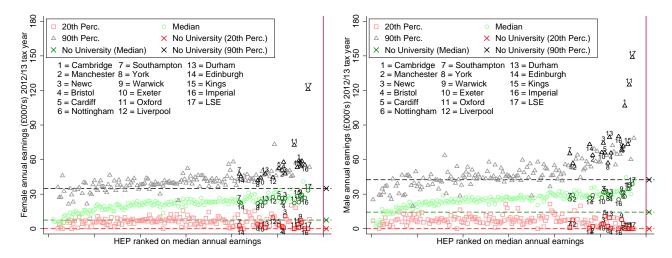


Figure 13: Predictive 20th, 50th and 90th percentile earnings for the 1999 cohort in 2012/13 for universities ranked on their graduates' 2012/13 median earnings. The predicted model uses an additive model based on the variables: the mean UCAS tariff of the institution, parental income, subject mix and region of institution. Non-university is included as a reference point, but note that these reflect unconditional earnings figures as the predictive model is not applied to this group. Note: The log scale is not used here. Ranking of institutions is based on estimated unconditional median earnings, not their predicted earnings.

## 5.3 Earnings Growth

Thus far we have focused on earnings variation at one point in time for a particular cohort. Our data also enable us to look at the development of graduates' earnings over time as they progress in their career. Figure 14 gives some insight into how the earnings of graduates from the 1999 cohort develop over the period 2008/09-2012/13, shown separately by gender (a similar figure for the 2002 cohort is in the appendix to this paper). We show earnings for all graduates, non graduates and graduates from our case study institutions. It is of course important to note that the period we consider here coincides with the 2008 economic recession and its aftermath, where across the

economy real earnings fell quite dramatically. This fall in real earnings particularly impacted the relatively young, such as those we are studying here. Thus this figure will be unlikely to be a representative period. The figure shows flat earnings with minimal growth for non graduates and female graduates but even in this period of economic turbulence, male graduates saw some growth and in particular male graduates from our case study institutions saw appreciable earnings growth. It appears that for males at least, attending a high status university did offer some protection from the impact of the recent recession.

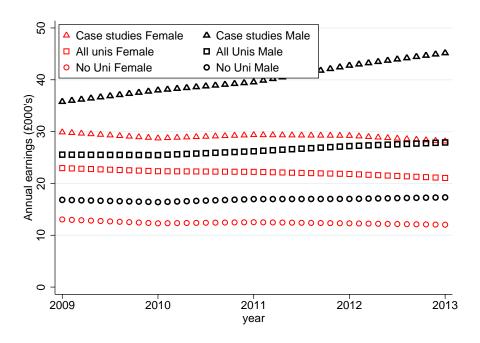


Figure 14: Mean earnings (including zeros) for a given cohort over time, by gender. This shows the 1999 cohort for 2008/09-2012/13.

# 6 Variation in earnings by subject and institution

We now turn to investigating the interaction between institution and subject to see if going to top institutions immunises the student against low earnings associated with certain subjects. We show that the answer varies by subject: there are subjects where institutions matter a great deal, others where it is not very relevant. We also show that for some institutions, subject choice really does matter, while for others, less so. In general, the choice of broader subject grouping (LEM, STEM or OTHER) appears to matter more than the choice between individual subjects.

### 6.1 Institutions and subject groups

We start by combining subjects into the aggregate subject groups described earlier, namely Law, Economics and Management (LEM); Science, Technology, Engineering and Mathematics (STEM) and OTHER (which predominantly consists of humanities based subjects). To consider whether some of the differences in earnings across subject group are largely attributable to differences in earnings across institutions, we plot the earnings of graduates for each institution at the 20th, 50th and 90th percentiles for each of these subject groups. We show this in a descriptive way, and the reader should note that these figures to not condition on background characteristics of the students.

Institutions are again ranked by median earnings, allowing us to compare graduates' earnings across the three subject groups within the same institution. Some institutions only have students in one subject group so only produce one observation - the clearest example of this is a group of small HEPs which specialize in humanities, which appear low in the ranked data and show only green dots.

If the different coloured shapes were on top of one another, that would imply that within a given institution there is little variation in graduates' earnings by subject group. However, we in fact observe something quite different; at the median and 90th percentile in most institutions LEM graduates have higher earnings than graduates in STEM or in OTHER subjects. This effect is stronger at institutions that have higher median graduate earnings, for which the separation is stark. The STEM results are somewhat higher than the OTHER category but the difference is much less clear. This therefore suggests that within institution, subject group choice is important.

The second question we are interested in is whether institution choice matters, conditional on subject choices. For both genders, there appears to be strong evidence that it does - LEM earnings are much greater for some institutions than others, for example - though at the 20th and 50th quantiles this evidence is less strong. At the 90th percentile, differences really do appear, suggesting that at the top end of the earnings distribution, institution choice matters more, though we again caution that we are not controlling for background characteristics of the students in these figures. Further, amongst the institutions that are highly ranked on median earnings, there are several which have very low earnings for OTHER subjects. Hence although institutional effects are large in these data, institution choice does not fully insure people against low earnings.

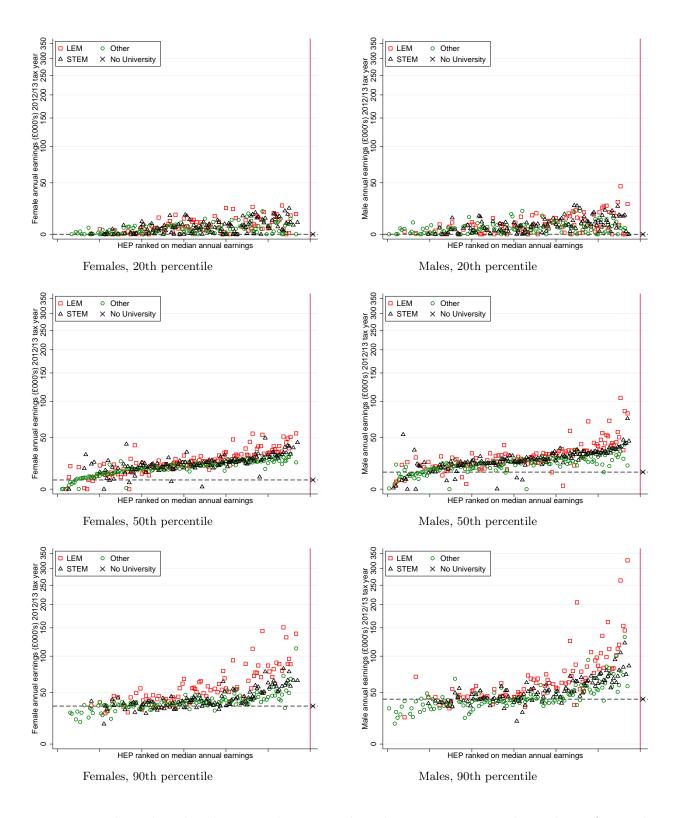


Figure 15: 20th, 50th and 90th percentile earnings by subject group. Note: the ranking of carried out by overall estimated median earnings of HEP, not by subject-specific median.

## 6.2 Institutions and individual subjects

We now continue to explore the interaction of institution and subject choice by investigating more detailed one digit JACS subject codes. Figure 16 shows a box plot of institutional median earnings by subject for women; Figure 17 shows the equivalent for men. Again, the reader should note that these plots do not condition on background characteristics of the students, meaning conclusions drawn here should be treated with caution.

To clarify what the Figures show, there is a maximum of one observation per institution per subject. An institution will have no observation for a given subject if it has no students (or only a handful of students) doing that subject and each institution-subject observation gets equal weighting in the plot. The aim of this is to describe the variation in earnings across institutions within subject, though the Figures are also revealing about the variation in earnings across subjects. At the bottom of the pictures the size of the black dots indicates the student numbers taking each subject. The line in the middle of each box is the median of institution medians. The top and bottom of the box are the 1st and 3rd quartile, so 75% of the institution medians fall within the rectangle. The whiskers which appear outside the rectangle show the rough scatter of the data, but some data can appear outside the whiskers (which would be the case if an institution had particularly high median earnings for a given subject).

To further illustrate the content of these Figures, the pink dot is the median earnings of graduates who attended the University of Southampton and who studied a Creative Arts degree. The point shows that even at a high status university Creative Arts students have relatively low median earnings, although we can see their institutional advantage from the fact that this median is higher than we see for most Creative Arts students who studied at other English HEPs.

From the Figures, it is clear that subject choice matters a lot in some cases, but much less so in others. Medicine and Economics stand out in particular in terms of their higher earnings (both are subjects with relatively few graduates), while graduates of Creative Arts and - to a lesser extent - Mass Communication tend to go on to achieve lower earnings. However, the differences in earnings amongst the remaining subjects are not so striking, suggesting the choice between those is less crucial.

From Figures 16 and 17 one can also observe that the importance of institution choice is also variable by subject. The range of median earnings for Maths and Computer Sciences, Engineering and Technology and Law is much greater than for Medicine, Biological Sciences and Mass Communication. This is particularly interesting for Medicine, as combined with the above point it shows that Medical graduates' earnings are relatively high, regardless of institution choice. Of course, the provision of Medicine in the UK is very tightly regulated so this result is not surprising, but it is

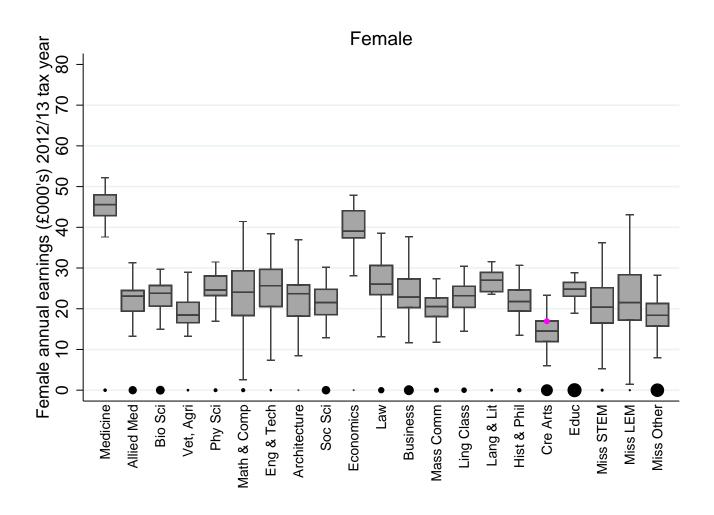


Figure 16: Box plots showing variation in institution-level median earnings at subject level for women for the 1999 cohort in 2012/13. Area of the blob size indicates number of students. The pink dot is for illustrative purposes, showing the median earnings of former graduates from the University of Southampton who studied Creative Arts.

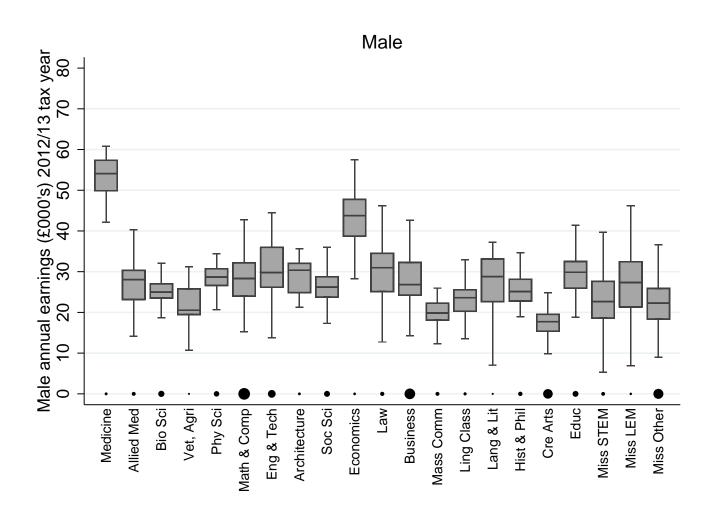


Figure 17: Box plots showing variation in institution-level median earnings at subject level for men for the 1999 cohort in 2012/13. Area of the blob size indicates number of students.

impactful. The same cannot be said for Economics graduates, for whom the returns appear to be high in expectation, but much less certain, while the opposite holds true for Creative arts and Mass Communication: both the level and variation in earnings for these subjects is low. In fact in terms of subsequent graduate earnings for both, even the relatively high achieving institutions (at the 75th percentile) are dominated by the relatively low achieving institutions (at the 25th percentile) for Medicine and Economics for both genders.

## 6.3 Institutions and individual subjects: model results

So far in this section we have investigated the importance of both subject choice conditional on institution and institution choice conditional on subject. However, thus far we have only displayed this without conditioning on important background characteristics. In this sub-subsection, we investigate these questions using more formal regression analysis.

Table 9 quantifies the differences in median graduate earnings for different subjects. The purpose of this table is to illustrate how median earnings vary by subject even after allowing for other sources of variation in earnings, namely student and institution characteristics. It relies on data from 2011/12 and 2012/13 for the 1999 cohort. Specifically, columns 1 and 2 shows the coefficients on subject dummy variables from a median regression of earnings with no additional controls (other than indicators of the year of the data) for females and males respectively. The results in these columns can therefore be interpreted as the raw or unconditional differences in earnings by subject, showing for example that the estimated median earnings for male (female) graduates in Medicine is around £51k (£41k) whilst for male (female) graduates in Creative Arts it is around £17k (£12k). The earnings variation across subjects is quite stark. To illustrate more concretely the variation in earnings that appears to be attributable to subject, in Columns 3 and 4 we present the earnings premiums for each subject relative to the earnings that graduates in the Creative Arts earn (Creative Arts was selected as it has a large student body and has low earnings). Hence male (female) graduates in Medicine have median earnings that are around £34k (£29k) greater than those with Creative Arts degrees.

Next we present results from the same model as previously but with additional controls for age, region, parental income and higher education fixed effects included. Again to illustrate the impact of these controls we show the earnings premiums for each subject compared to the earnings for Creative Arts. After these controls are added the premiums associated with Medicine for males (females) is substantially lower at £24k (£23k). Hence some of the higher premiums that Medicine attracts is attributable to other factors, including the fact that Medicine is studied at high tariff institutions. Nonetheless even after allowing for these factors, medical graduates earn significantly more than those from other subjects.

After allowing for these characteristics, the earnings premiums associated with some subjects is reduced. For instance, the premium associated with Economics for males (females) is reduced from £23k (£20k) to £13k (£18k). Again this indicates that the high premium for Economics is partly attributable to other factors, including the institutions that offer this subject. This specification does not however control for the student intake in that particular subject at that specific institution, which may differ from the average student intake at that institution as a whole. In the final two columns we control for the subject/institution characteristics using HESA data, including the tariff points of students in that subject/institution combination. We have to remove institution fixed effects to enable the model to converge. The pattern of results does not change markedly though the coefficients on medicine are reduced somewhat and some subject differences that were previously statistically insignificant become marginally so and vice versa. In summary, the differences in earnings across subjects get compressed once we take account of the fact that graduates with different characteristics take different degree subjects at different institutions. Therefore, although part of the reason why Creative Arts graduates have very low earnings is because they possess characteristics that would be associated with lower earnings anyway, this does not explain away all of the earnings differences. This suggests that for a given individual it still holds true that the subject choice between creative arts and the rest remains important.

Subject	Subject Group	Raw earnings (£000's)		Raw differences from Cre Arts (£000's)		Conditional differences from Cre Arts HEP f.e. (£000's)		Conditional differences from Cre Arts HESA (£000's)	
		M	F	M	F	M	F	M	F
Medicine	STEM	50.7	41.1	34.1***	28.6***	24.1***	23.2***	21.0***	21.8***
Economics	LEM	$\frac{2.8}{39.9}$	32.6	23.3***	5.6 20.1***	$\begin{array}{c} 8.3 \\ 13.2 \end{array}$	7.8 17.8**	4.2 13.9***	20.0
Economics	1717171	39.9	32.0	5.5	4.2	8.8	7.2	4.8	20.0 16.3
Eng & Tech	STEM	29.1	19.8	12.5***	7.4*	6.4	5.1	7.9*	6.1
<u> </u>		3.2	3.1	4.5	4.3	8.1	6.6	4.1	4.9
Educ	Other	28.0	21.3	11.4**	8.9**	9.1	7.9	10.0**	8.7*
т	TEM	3.4	3.1	4.6	4.4	8.2	7.2	4.0	4.8
Law	LEM	28.0	23.3 $3.0$	11.3**	10.9**	7.1	8.7 7.0	7.1*	7.8*
Phy Sci	STEM	27.7	21.2	11.0**	4.3 $8.7*$	3.5	6.4	4.1	6.0
<i>y</i>		3.1	3.7	4.4	4.8	8.2	7.5	3.9	4.9
Architecture	Other	26.4	17.6	9.8**	5.1	6.4	5.2	6.8	3.3
		3.1	3.8	4.4	4.9	8.1	7.2	4.7	4.7
Allied Med	STEM	26.4	21.3	9.8**	8.9**	6.0	7.3	6.6*	7.7
Lang & Lit	Other	$\frac{3.2}{26.3}$	$\overset{\it 3.0}{23.8}$	9.7**	4.3 11.3**	$\frac{8.0}{2.5}$	$\frac{7.0}{4.8}$	3.9 1.3	4.7 $4.7$
Lang & Lit	Other	3.2	<b>4.</b> 0	4.5	5.1	8.2	7.6	3.9	6.4
Math & Comp	STEM	25.9	20.4	9.2**	8.0*	6.3	6.7	7.1*	7.0
		3.2	2.8	4.5	4.2	8.1	6.7	4.0	4.7
Miss LEM	LEM	25.8	18.5	9.2*	6.1	8.3	6.2	6.6	6.1
D	LEM	3.5	3.1	4.7	4.3	8.1	7.3	4.2	4.9 C. 1
Business	LEM	24.9	20.0 3.2	8.3*	7.5*	5.9 8.1	6.6 7.2	6.1	6.1 4.7
Hist & Phil	Other	24.6	20.8	7.9**	$\frac{4.4}{8.4^*}$	1.6	4.5	2.1	3.3
		2.3	3.6	3.9	4.7	8.2	7.4	3.8	4.8
Soc Sci	Other	24.5	18.4	7.9*	6.0	4.2	4.9	4.5	5.2
D: 0.	OPPD 6	3.3	2.9	4.6	4.2	8.2	6.9	4.0	4.7
Bio Sci	STEM	22.8	21.9	6.1	9.5**	1.1	6.0	2.0	5.9
Miss STEM	STEM	3.2 21.9	18.3	$\frac{4.5}{5.3}$	$\begin{array}{c} 4.6 \\ 5.8 \end{array}$	8.2 4.2	$\frac{7.4}{5.0}$	4.1	$\frac{4.8}{3.9}$
111100 0 1 12111	012111	3.6	3.5	4.7	4.6	8.1	7.0	4.1	5.1
Ling Class	Other	21.6	20.8	4.9	8.3*	9	4.8	1	4.6
		3.5	3.1	4.7	4.4	8.5	7.0	4.6	4.7
Miss Other	Other	21.2	17.3	4.6	4.9	4.7	5.6	4.5	5.0
Mass Comm	Other	$\frac{3.2}{20.9}$	$\frac{3.0}{15.4}$	$\frac{4.5}{4.2}$	$\frac{4.3}{2.9}$	3.7	$\frac{7.0}{2.8}$	3.5	$\frac{4.8}{2.0}$
Wass Collini	Other	3.9	3.0	5.0	4.3	8.1	7.0	4.2	4.7
Vet, Agri	STEM	20.1	18.2	3.5	5.8	4.3	4.8	4.2	5.5
, 0		4.0	3.3	5.1	4.5	8.1	7.3	4.9	4.9
Cre Arts	Other	16.7	12.4						
37		3.1	3.0	3.7	3.7	3.7	3.7	3.7	3.7
Years Regions		Yes No	Yes No	Yes No	Yes	Yes Yes	$\begin{array}{c} { m Yes} \\ { m Yes} \end{array}$	Yes Yes	$\begin{array}{c} { m Yes} \\ { m Yes} \end{array}$
Age		No	No	No No	No No	Yes	Yes	Yes	Yes
Rich		No	No	No	No	Yes	Yes	Yes	Yes
HEI		No	No	No	No	Yes	Yes	No	No
HESA		No	No	No	No	No	No	Yes	Yes
N		21,184	24,058	21,184	24,058	21,184	24,058	21,184	24,058

Table 9: Subject median regressions. The first two columns of results show raw estimated earnings by subject. The next two show the difference in earnings from Creative Arts. The next two columns show the conditional differences from Creative Arts - i.e. the differences once controls for region, age, parental income and HEI fixed effects are included. The final two columns show differences from Creative Arts with controls for region, age, parental income and HESA data included. All figures are in £000's. Uses 2011/12 and 2012/13 data and the 1999 cohort (estimates are given for 2012/13). Standard errors are clustered at HEP level. \* indicates significantly different to the base (CreArts) at 10% level, \*\* 5% and \*\*\* 1%.

Whilst Table 9 quantifies the differences in graduates' earnings associated with different degree subjects, Table 10 shows the differences in median graduate earnings associated with different types of institution. The purpose of this table is to illustrate how median earnings vary by institution group (categorised by decile of mean UCAS tariff), even after allowing for other sources of differences in earnings, namely student characteristics and subject of degree. The tables use data from 2011/12 and 2012/13, and presents estimated earnings for the 1999 cohort specifically.

Columns 1 and 2 provide estimates of median graduate earnings by institution group for females and males respectively, from a median regression which controls only for year of observation. The variation in median earnings across institution group is sizeable, with estimated male median earnings ranging from around £20k for those in institutions in the lowest two decile groups through to nearly £40k for those in the top 5 institutions. For women the variation is somewhat lower, ranging from £15k for those in institutions in the bottom two decile groups through to £26k for those in the top decile group. Further, for males there is a sizeable gap in earnings between those in the top 5 institutions and those in the rest of the top decile group. For women, the earnings of those in the top 5 institutions are not any higher than those in the rest of the top decile group.

Columns 3 and 4 follow the pattern from Table 9 and present the earnings premiums for females and males respectively compared to the earnings of graduates in group G0. The results indicate that median male (female) earnings in the top 5 institutions is around £18k (£11k) more than those in the bottom two decile institution groups.

In the columns 4 and 5 the model includes controls for demographic factors (cohort, region, age, rich, and subject fixed effects) but does not take account of the characteristics of the student intake from the HESA data. In columns 4 and 5 it is apparent that much of the variation that appears to be across institutions can be attributable to other factors. The median male (female) earnings premium associated with being in a top 5 institution is reduced to £14k (£9k) compared to the earnings of graduates in the bottom two decile institution groups. The final two columns then add in controls for the HESA characteristics, including the mean tariff of the student intake. The coefficients remain similar in magnitude in most cases, suggesting a premium for those attending institutions in the top groups but the magnitude of the standard errors increases dramatically and results become statistically insignificant. We conclude that there is clearly less variation in graduate median earnings by institution group once one takes account of student characteristics and degree subject. However for males, institution of study may be potentially an important determinant of median earnings for those in the top groups but the methodology and quantity of data we use in this model does not enable us to give sufficient precision to these estimates.

HEP	Raw		Raw		Conditional		Conditional		
Group	earnings		differences		differ	differences		differences	
	(£000's)		fron	from G0		from G0		from G0	
		'		£000's		Subject		HESA	
						(£000's)		(£000's)	
	M	F	M	F	M	F			
G10high	38.7	26.8	17.9***	10.6***	13.5***	8.4***	13.0	3.2	
	1.1	1.0	3.0	2.9	3.2	1.8	9.9	10.7	
G10low	33.0	26.1	12.2***	9.8***	9.5***	7.3***	11.6	4.7	
	2.0	1.8	3.4	3.3	3.4	2.2	9.7	10.8	
G9	31.4	25.5	10.6***	9.3***	7.2**	7.4***	9.1	5.6	
	1.5	1.2	3.2	3.0	3.2	1.7	9.7	10.6	
G8	29.3	23.5	8.4***	7.3**	5.9*	6.4***	8.3	6.2	
	1.5	1.3	3.2	3.0	3.3	1.8	9.8	10.6	
G7	28.5	21.3	7.7**	5.0	5.5*	4.5**	7.9	5.8	
	1.7	1.7	3.3	3.2	3.2	2.1	9.8	10.7	
G6	24.7	19.7	3.8	3.5	3.8	4.1**	6.6	6.2	
	1.9	1.5	3.4	3.1	3.2	1.9	9.8	10.6	
G5	23.8	19.9	3.0	3.7	2.8	3.8**	5.6	6.0	
	1.6	1.3	3.2	3.0	3.3	1.7	9.9	10.4	
G4	22.2	17.6	1.4	1.3	1.7	2.9*	6.0	6.0	
	1.4	1.2	3.2	3.0	3.2	1.7	9.9	10.5	
G3	22.8	19.6	2.0	3.3	1.4	4.1**	5.5	7.4	
	1.3	1.4	3.1	3.1	3.2	2.0	10.0	10.6	
G2	24.2	18.4	3.4	2.2	2.3	2.1	5.9	5.9	
	1.3	1.2	3.1	3.0	3.2	1.7	10.0	10.4	
G1	19.2	14.7	-1.6	-1.6	6	.2	5.2	5.7	
	1.3	1.1	3.1	2.9	3.3	1.8	10.1	10.5	
G0	20.8	16.3							
	2.0	1.9							
Years	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	
Regions	No	No	No	No	Yes	Yes	Yes	Yes	
Age	No	No	No	No	Yes	Yes	Yes	Yes	
Rich	No	No	No	No	Yes	Yes	Yes	Yes	
Subject	No	No	No	No	Yes	Yes	Yes	Yes	
HESA	No	No	No	No	No	No	Yes	Yes	
N	21,184	$24,\!058$	21,184	24,058	21,184	24,058	21,184	24,058	

Table 10: University median regressions. The first two columns of results show raw estimated earnings by HEP group. The next two show the difference in earnings from Group 0. The next two columns show the conditional differences from Group 0 - i.e. the differences once controls for region, age, parental income and subject are included. The final two columns show the conditional differences controlling for HESA characteristics as well. All figures are in £000's. Uses 2011/12 and 2012/13 data and the 1999 cohort (estimates are given for 2012/13). Standard errors are clustered at HEP level. \* indicates significantly different to the base (G0) at 10% level, \*\* 5% and \*\*\* 1%.

In summary, we find a significant amount of variation in graduates' earnings by both subject and institution. For males, even after allowing for other factors that influence earnings, variation in earnings by subject and potentially institution is quantitatively important. For women, variation by institution is less marked. What is also notable however, is that some subjects attract particularly high earnings and hence these particular subject choices will have a stronger association with graduates' earnings than choice of institution. For example, for males the decision to study Medicine is more important in terms of expected median earnings than say an individual choosing an institution in the top decile group instead of one in the bottom decile group. Of course Medicine is a very competitive occupation and selection into that subject is likely to be on the basis

of unobserved characteristics, such as IQ, motivation and familiarity with the medical field. We cannot therefore take the differences in earnings presented in Table 9 as being causal. In any case, Medicine and, for males, Economics, are outlier subjects attracting particularly high earnings. In the case of males, there is very little variation in median earnings across many subjects (including engineering, education, Law, physical sciences, architecture, subjects allied to Medicine, Maths, computing, Business studies and social science) once we account for student and institutional characteristics. Similarly for women, Medicine is a higher earning subject and the decision to take this subject has a stronger association with earnings than institution choice. However, for women across a range of other subjects there is little variation in median earnings (for example across Law, physical sciences, subjects allied to Medicine, Maths, computing, Business studies, social science and biological science), again once we account for student and institutional characteristics. This finding aligns with the earlier unconditional figures and suggests the key differences by institution are at the top end of the distribution rather than at the median of graduate earnings.

## 7 Variation in earnings by parental income

We now turn our focus to the variation in earnings that we observe across graduates from different family backgrounds, using the measure of parental income we introduced in Section 2. We reiterate that our measure of family income is a relatively crude proxy and again note that we are unable to control for individuals' own prior achievement before they enter higher education, meaning these results must be interpreted cautiously. Nonetheless understanding how earnings vary by family background, and particularly whether graduates from poorer family backgrounds go on to have more earnings variability, is important.

Figure 18 shows the earnings distribution for graduates from higher income households (green triangles), graduates from lower income households (red circles) and for non graduates (grey crosses) for the 1999 cohort in 2012/13, by gender. Points to the right of each figure show the mean for each group. The results are striking; graduates from higher income households earn more right across the distribution, from the 20th percentile upwards, for both females and males. Whilst graduates from both lower and higher income households earn more than non graduates, the gap between the lower and higher income groups is sizeable, particularly at the very top of the distribution. One of our research questions was whether there is more variation in the earnings of graduates from lower income households. In fact there is less variation in earnings within the group of graduates from lower income households but this is largely driven by the very high earnings of the top earners from higher income households.

There are many possible reasons for the divergence in earnings between graduates from lower

and higher income households. One such reason might be that graduates from more modest family backgrounds access higher status universities for example. Figure 19 takes the first step to addressing this possibility by plotting median earnings for graduates from higher income (green triangles) and lower income (red squares) households for each institution in 2012/13. Institutions are ranked left to right in terms of overall median earnings. To the right of the figure the median for non-university is indicated. Even within institutions the median earnings of graduates from higher income households tend to be above the median earnings for those from lower income households, suggesting that broadly speaking even when comparing graduates from the same institution, those from a higher income background go on to do better in terms of labour market earnings. This is concerning as graduates within the same institution should be relatively comparable in the UK's highly stratified HEPs.

However, it still may be the case that other characteristics of graduates from higher and lower income households differ and that this may explain some of the patterns we are seeing. For example, the gap could in earnings could be driven by graduates from higher income households choosing subjects that provide an easier route to higher earnings.

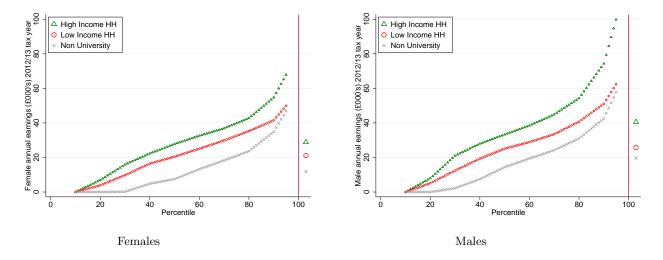


Figure 18: Earnings distribution for individuals in the 1999 cohort in 2012/13 for those from higher income households vs. individuals from lower income households. Note: The log scale is not used here.

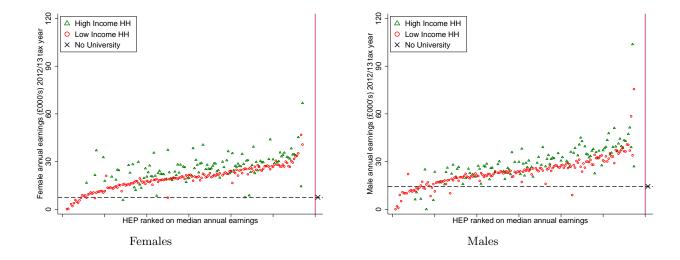


Figure 19: Unconditional median earnings for graduates from higher and lower income households for universities ranked on their graduates' 2012/13 median earnings. Note: The log scale is not used here.

To explore this more formally we turn to table 11, which shows estimated differences in earnings (at the 20th, 50th and 90th percentiles) for students from higher income backgrounds compared to those who come from relatively lower income households. The table uses data from 2011/12 and 2012/13 for the 1999 cohort. There are three horizontal panels: the top panel shows estimated earnings differences between graduates from higher and lower income households at the 20th percentile (from a quantile regression). The middle panel shows differences between these groups at the 50th percentile and the bottom panel at the 90th percentile. In the first two columns of each panel, the upper numbers are the earnings of female and male graduates respectively from higher income families, the middle numbers the earnings of those from relatively lower income families and the bottom numbers the percentage difference between the two, expressed as a premium for individuals from high income households over those from low income households. In the first two columns the model only includes controls for year of observation. Columns 3 and 4 then show the same model, but this time expressed as a raw earnings difference from the earnings of graduates from lower income families (which are therefore entered as zero in these columns). Columns 5 and 6 then show the results using the same format but from a model which includes a set of controls, namely age, region, subject of degree and student characteristics from the HESA data.

The unconditional results in columns 1-4 show that the earnings of graduates coming from higher income households are higher than the earnings of graduates from lower income households, right through the distribution. This premium is sizeable and shows an interesting U shape for both genders; it is 61% at the 20th percentile, 30% at the 50th percentile and 58% at the 90th percentile for men, while the equivalent figures for women are 46%, 24% and 33%. In terms of raw earnings

differences, these figures equate to £8k, £8k and £31k for men and £4k, £5k and £13k for women.

However, of course some of the apparent premium earned by those from a higher income household may actually be attributable to other characteristics of students and the degrees they study. Columns 5 and 6 show that when conditioning variables are included, the premium associated with coming from a higher income background does indeed fall, to 16% at the 20th percentile, 11% at the median and 20% at the 90th percentile for males, while the equivalent premiums for females are reduced to 13%, 9% and 14%. These results remain statistically significant at the 1% level, and are sizeable gaps in earnings when one considers that these are between graduates with similar characteristics, taking the same subject and attending similar institutions. Higher education does not therefore appear to have eliminated differences in earnings between students from lower and higher income backgrounds.

Further it is interesting that the U shape relationship is preserved in columns 5 and 6. This shows that while the impact of coming from a high income background is strong right through the distribution, in particular it helps protect against low earnings, and provides much greater opportunity for much higher earnings.

We reiterate that our approach here does not allow us to necessarily assign causality to these relationships, due to unobservable characteristics we are unable to control for, such as intelligence or degree classification. However, on the other hand, we believe our crude measure of parental income almost certainly biases the impact down. We only include those who borrow the maximum amount available to rich individuals, and therefore miss individuals from high income households who borrow less than this, and also incorrectly identify those from low income households who borrow at the rich maximum. Further, individuals who do not borrow at all are completely disregarded from this analysis. This is also likely to understate the overall impact, as one might expect these individuals to both be wealthy and go on to have high earnings (for example we know that privately educated individuals have a significant wage premium).

	R	aw	Ra	aw	Condi	tional		
	earr	nings	differences		differences			
	(£000's)		from	from low		from low		
	(**************************************		family income		family income			
			(£00	00's)	(£000's)			
	M	$\mathbf{F}$	M	F	M	F		
			20th Pe	ercentile				
Higher family income earnings	20.3	14.0	7.7***	4.4***	3.0***	1.7***		
	1.1	.7	.9	.6	.7	.6		
Lower family income earnings	12.6	9.6						
	.5	.3						
% Wage Premium	61.1	45.8	61.1	45.8	16.1	13.4		
			50th Pe	50th Percentile				
Higher family income earnings	35.0	27.8	8.0***	5.3***	3.3***	2.1***		
	1.0	.7	.9	.6	.6	.6		
Lower family income earnings	27.0	22.5						
	.5	.4						
% Wage Premium	29.6	23.6	29.6	23.6	10.9	8.6		
		90th Percentile						
Higher family income earnings	84.0	54.9	30.8***	13.5***	10.7***	6.2***		
	7.0	2.1	6.8	2.0	1.8	1.3		
Lower family income earnings	53.2	41.4						
	1.5	.8						
% Wage Premium	57.9	32.6	57.9	32.6	19.6	14.3		
Years	Yes	Yes	Yes	Yes	Yes	Yes		
Regions	No	No	No	No	Yes	Yes		
Age	No	No	No	No	Yes	Yes		
Subject	No	No	No	No	Yes	Yes		
HEI	No	No	No	No	No	No		
HESA	No	No	No	No	Yes	Yes		
N	18,038	20,413	18,038	20,413	18,038	20,413		

Table 11: Earnings differences for graduates from lower and higher income households at the 20th, 50th and 90th percentiles estimated from quantile regression models. Note that zero earnings are excluded from these regressions. High family income premium indicates the additional earnings for graduates from a higher income household. Low family income earnings indicates earnings of graduates from a lower income background. Percentage wage premium calculates the wage premium for those coming from a richer family background compared to the earnings of those from lower income households, assuming all controls are held constant across the two groups at their means. The first two columns of results show raw estimated earnings for high and low household income earnings. The next two show the difference in earnings from low household income. The final two columns show the conditional difference from low household income - i.e. the difference once controls for region, age, subject and student characteristics are included. All figures are in £000's. Uses 2011/12 and 2012/13 data and the 1999 cohort (estimates are given for 2012/13). Standard errors are clustered at HEP level. \* indicates significantly different to the base (lower family income) at 10% level, \*\* 5% and \*\*\* 1%.

# 8 Conclusions and policy implications

Using an innovative administrative data set, consisting of hard linked HMRC and SLC individual level data and HESA aggregates, we document the earnings of graduates, focusing particularly on variation across subject, institution and an indicator of student family income and building on previous work in Britton et al. (2015). The paper is the first of its kind to use such data in the English context, one of very few studies internationally that has used administrative data to examine issues relating to social mobility, and the first to examine the correlation between a measure of parental income and graduates' income whilst being able to take account in some detail of the

type of higher education attended. What is clear from the data is the sheer scale of the variation in graduate earnings, even between graduates from the same institutions and taking the same subjects. Just to take one example, male graduate earnings from the LSE range from £170k at the 90th percentile to nearer £40k at the median. Whilst 15% of graduates have zero earnings, at the other end of the spectrum graduates from a wide range of institutions have very high earnings. For example, in the case study institutions that we studied, graduates from all our named institutions earn £60k or above at the 90th percentile of the distribution. Even at institutions with far lower median earnings, graduates at the 90th percentile earn £30k or more. There is no doubt that a degree offers a pathway to relatively high earnings for a large subset of graduates, from across a range of institutions.

Further these data illustrate the substantial earnings advantage of graduates as compared to non-graduates, not least because only 15% of the former have zero earnings compared to 30% of the latter. Indeed, graduates from almost all universities earn more than individuals at the 20th percentile of the non-university earnings distribution.

Throughout the paper we refer to graduates, though in practice we are investigating the earnings of borrowers, as we have no information on whether individuals complete their course. This is an important distinction to the extent that borrowers who do not complete their degrees may go on to have lower earnings. On the other hand, we also do not observe graduates who did not borrow. Given that these individuals are more likely to be from higher income households, and we know that those from high income households are more likely to subsequently have high earnings, this may underestimate the earnings of graduates. However, since only around 10% of borrowers do not complete their course and only around 15% of students don't borrow, we do not believe our overall findings in this paper will be dramatically impacted. Further, in terms of thinking about the long run cost to government of issuing student loans, this is precisely the population of interest.

Though we find significant variation in graduate earnings by subject and institution, much of the variation is actually attributable to differences in the characteristics of students taking different degree options. Certainly there are some degree subjects with very high earnings like Medicine and Economics, irrespective of institution attended and the characteristics of students who study the subject. Equally there are some subjects that are associated with very low subsequent earnings, even allowing for student characteristics and institution. The clearest example of this is Creative Arts. Over the period we observe in our data, the proportion of students taking subjects such as Economics, Law and Maths and Computer Sciences has reduced marginally, and more take subjects such as Creative Arts. It is too early to determine whether this is a trend, although Universities UK data do suggest stronger growth between 2002 and 2011 in business & administrative studies,

biological sciences, education, social studies and creative arts & design and rather weaker growth in Law, Mathematical Sciences and Computer science, to name but a few. Graduates who study the creative arts, for example, tend to earn less and so over time we might be concerned that these shifts may bring down the aggregate graduate earnings premium. What is not clear is the reason for these changes in subject mix. It may be that students prefer some lower earning subjects, irrespective of expected earnings. It may however be a supply side issue, with institutions preferring to offer more places for lower cost courses since fees do not typically vary by subject. Staffing creative arts degrees is likely to be much cheaper than staffing degrees in Economics, Law and Maths and Computer Science. These findings have implications for our understanding of the nature of subsidy of higher education. Given the relatively low earnings of graduates with degrees in some subjects, the level of public subsidy for these graduates is likely to be greater than for other graduates in other subjects, such as economics, even given the lower costs of provision for some subjects as compared to others. Making this explicit when considering the shape of higher education and in particular where any further expansion might take place would seem important.

One purpose of this paper is to provide a proof of concept for using the data to provide some useful information that might inform students' choice of degree. However, although we are able to document the variation in earnings across graduates from different institutions, once we start studying subject-institution combinations we find that the 10% sample of data that we are using from HMRC is still not large enough to look in detail at large numbers of higher education institution and subject combinations without making strong econometric modelling assumptions. This limits our ability to provide sufficient information about every subject-institution combination and hence the usability of the database, as it stands, for information provision to students. However, there is an easy solution to this problem, which is to utilise the full database instead of the 10% sample that has so far been made available to us. This larger database should provide the scale needed for the data to be sufficiently granular to provide information to students.

A main finding from this paper is that graduates' family background - specifically whether they come from a lower or higher income household - continues to influence graduates' earnings long after graduation. Graduates from higher income households earn more (up to around 60% more for males and 45% for females) than their peers from lower income households. This gap is by no means entirely explained by differences in the subjects studied or institutions attended by graduates from higher or lower income households, though it is substantially reduced once we account for these factors. When we take account of different student characteristics, degree subject and institution attended, the gap between graduates from higher and lower income households is still a sizeable, at around 10% at the median. Further, we find that the gap is larger at the

20th and 90th percentiles of the graduate earnings distribution, suggesting coming from higher income households both protects against low earnings and provides greater opportunity for very high earnings. The magnitude of this effect is sufficient to be important. Not only does it hold when we condition on subject choice and a wide range of institutional characteristics, but also it holds despite the large amount of measurement error in our relatively crude measure of parental income, that will almost certainly bias our estimates towards zero. This finding raises questions about the extent to which higher education can ensure that the labour market prospects of students from lower and higher income backgrounds are similar.

There are several possible explanations for our results. Students from wealthier families may have greater financial support from parents, may be more likely to relocate for work and may also be able to take greater career and financial risks than students from poorer backgrounds. They might have access to financial, social and cultural capital (e.g. networks). Alternatively, graduates' from lower income backgrounds experience overt or covert discrimination in the labour market that constrains their earnings. All these factors will influence their career prospects and earnings and may explain why there continues to be a gap in the earnings of students from rich and poor backgrounds even after they experience the same higher education.

An additional explanation for our results is that there are some unobserved characteristics of the students from higher income households that are correlated with their family income and also their own subsequent earnings. For example, differences in graduates' ability (IQ, degree class), social skills or determination, to name but a few. Indeed in our data we are only able to control for the average level of prior achievement across all individuals taking a particular subject-institution combination, rather than the individual's own level of achievement at A level. This means we are not able to discount the possibility of ability bias.

Whatever the explanation, there is a need for further investigation into this issue, not least to inform policy and practice of universities that may play a role in assisting students, particularly those from poor backgrounds, to make the transition into the labour market. One step that would be particularly helpful would be to link HMRC data to the National Pupil Database and data from the Higher Education Statistics Agency to enable us to compare students with identical school achievement who come from higher/lower income households and reduce ability bias. With this additional data will we be able to estimate models that better control for the individual's own level of pre higher education achievement.

## References

- Arcidiacono, P. (2004). Ability sorting and the returns to college major. *Journal of Econometrics* 121(1), 343–375.
- Barr, N. (2007). Financing higher education: tax, graduate tax or loans? In J. L. G. Hills, John and D. Piachaud (Eds.), *Making Social Policy Work: Essays in honour of Howard Glennerster*, pp. 109–130. Bristol: Policy Press.
- Barr, N. and N. Shephard (2010). Towards setting student numbers free. Unpublished paper: London School of Economics.
- Becker, G. S. (1962). Investment in human capital: A theoretical analysis. *The Journal of Political Economy*, 9–49.
- Bhuller, M., M. Mogstad, and K. G. Salvanes (2011). Life-cycle bias and the returns to schooling in current and lifetime earnings. *NHH Dept. of Economics Discussion Paper* 4.
- Black, S. E., P. J. Devereux, and K. G. Salvanes (2005). The more the merrier? the effect of family size and birth order on children's education. *The Quarterly Journal of Economics*, 669–700.
- Blanden, J., P. Gregg, and L. Macmillan (2007). Accounting for intergenerational income persistence: Noncognitive skills, ability and education. *Economic Journal* 117, C43–C60.
- Blundell, R., L. L Dearden, and B. Sianesi (2005). Evaluating the effect of education on earnings: models, methods and results from the national child development survey. *Journal of the Royal Statistical Society, Series A* 168, 473–513.
- Bratti, M., R. Naylor, and J. Smith (2005). Variations in the wage returns to a first degree: Evidence from the british cohort study 1970.
- Britton, J., N. Shephard, and A. Vignoles (2015). Comparing sample survey measures of english earnings of graduates with administrative data. Unpublished paper: Department of Economics, Harvard University.
- Browne, J. (2010). Securing a sustainable future for higher education: an independent review of higher education funding and student finance.
- Bukodi, E. and J. Goldthorpe (2011a). Social class returns to higher education: Chances of access to the professional and managerial salaries for men in three British birth cohorts. *Longitudinal and Life Course Studies* 2, 1–71.

- Bukodi, E. and J. H. Goldthorpe (2011b). Class origins, education and occupational attainment in britain. *European Societies* 13(3), 347–375.
- Card, D. (1999). The causal effect of education on earnings. In *Handbook of labor economics*, Volume 3, pp. 1801–1863.
- Card, D. (2012). Introduction to earnings, schooling, and ability revisited. 35th Anniversary Retrospective (Research in Labor Economics, Volume 35) Emerald Group Publishing Limited 35, 107–110.
- Card, D., R. Chetty, M. Feldstein, and E. Saez (2010). Expanding access to administrative data for research in the United States. Unpublished paper: Department of Economics, Harvard University.
- Carneiro, Pedro, I., L. Garcia, K. G. Salvanes, and E. Tominey (2013). Intergenerational mobility and the timing of parental income. *In CES Ifo Conference on Economics of Education*.
- Chevalier, A. (2011). Subject choice and earnings of UK graduates. *Economics of Education Review 30*, 1187–1201.
- Chowdry, H., C. Crawford, L. Dearden, A. Goodman, and A. Vignoles (2013). Widening participation in higher education: analysis using linked administrative data. *Journal of the Royal Statistical Society: Series A (Statistics in Society)* 176(2), 431–457.
- Chowdry, H., L. Dearden, A. Goodman, and W. Jin (2012). The distributional impact of the 2012-13 Higher Education funding reforms in England. *Fiscal Studies* 33, 211–236.
- Crawford, C. and A. Vignoles (2014). Heterogeneity in graduate earnings by socio-economic background. Unpublished paper: Institute of Fiscal Studies.
- Croxford, L. and D. Raffe (2013). Differentiation and social segregation of uk higher education, 1996 2010. Oxford Review of Education 79, 172–192.
- Dolton, P. and A. Vignoles (2000). The incidence and effects of overeducation in the uk graduate labour market. *Economics of Education Review 19*, 179–198.
- Dorling, D. (2016). A better politics: How government can make us happier. *London Publishing Partnership*.
- Figlio, D. N., K. Karbownik, and K. G. Salvanes (2015). Education research and administrative data. *National Bureau of Economic Research No. w21592*.

- Green, F., S. Machin, R. Murphy, and Y. Zhu (2012). The changing economic advantage from private schools. *Economica* 79, 658–679.
- Hussain, I., S. McNally, and S. Telhaj (2009). University quality and graduate wages in the UK. CEE Discussion Paper No. 99.
- Ladd, H. and R. Walsh (2002). Implementing value-added measures of school performance: getting the incentives right. *Economics of Education Review 21*, 1–18.
- Macmillan, L., C. Tyler, and A. Vignoles (2013). Who gets the top jobs? the role of family background and networks in recent graduates access to high-status professions. *Journal of Social Policy*, 1–29.
- Monks, J. (2000). The returns to individual and college characteristics: Evidence from the national longitudinal survey of youth. *Economics of Education Review 19*, 279–289.
- Naylor, R. (2002). Sheer class? the extent and sources of variation in the UK graduate earnings premium. LSE STICERD Research Paper No. CASE054.
- Savage, M. and R. Burrows (2009). Some further reflections on the coming crisis of empirical sociology. *Sociology* 43, 762–772.
- Sloane, P. J. and N. C. O'Leary (2005). The return to a university education in Great Britain.

  National Institute Economic Review, 75–89.
- Smith, J. and R. A. Naylor (2001). Determinants of individual degree performance: Evidence for the 1993 UK university graduate population from the USR. Oxford Bulletin of Economics and Statistics 63, 29–60.
- Walker, I. and Y. Zhu (2011). Differences by degree: evidence of the net financial rates of return to undergraduate study for England and Wales. *Economics of Education Review 30*, 1177–1186.
- Walker, I. and Y. Zhu (2013). The impact of university degrees on the lifecycle of earnings: some further analysis. *BIS Research Paper No.112*.
- Webber, R. (2009). Response to 'The coming crisis of empirical sociology: an outline of the research potential of administrative and transactional data'. Sociology 43, 169–178.

# 9 Appendix

## 9.1 Appendix A

### 19 JACS codes

- A Medicine and Dentistry
- B Subjects allied to Medicine
- C Biological Sciences
- D Veterinary Sciences, Agriculture and related subjects
- F Physical Sciences
- G Mathematical and Computer Sciences
- H Engineering
- J Technologies
- K Architecture, Building and Planning
- L Social studies
- M Law
- N Business and Administrative studies
- P Mass Communications and Documentation
- Q Linguistics, Classics and related subjects
- R European Languages, Literature and related subjects
- T Eastern, Asiatic, African, American and Australasian Languages, Literature and related subjects
  - V Historical and Philosophical studies
  - W Creative Arts and Design
  - X Education

## 9.2 Additional figures and tables

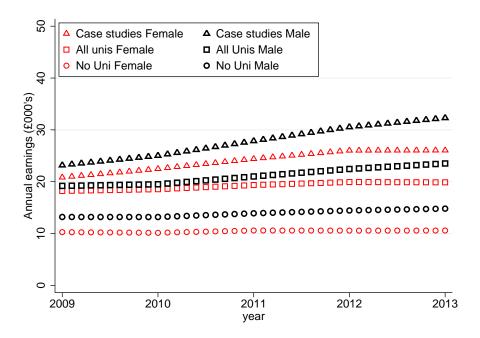


Figure 20: Mean earnings (including zeros) for a given cohort over time, by gender. This shows the 2002 cohort for 2008/09-2012/13.

							$\% \mathrm{Own}$
	All	$\mathbf{M}$	$\mathbf{F}$	% Pop	% F	%G10	Region
East Midlands	1,720	745	975	8.0	56.7	10.6	45.9
East	2,179	986	1,193	10.1	54.7	12.9	20.7
London	3,485	1,590	1,895	16.1	54.4	8.2	58.6
North East	956	403	553	4.4	57.8	9.2	71.8
North West	3,189	1,457	1,732	14.7	54.3	7.4	66.0
South East	3,791	1,802	1,989	17.5	52.5	11.1	38.1
South West	2,125	978	1,147	9.8	54.0	8.5	44.5
West Midlands	2,296	1,045	1,251	10.6	54.5	9.8	49.3
Yorks & the Humber	1,881	862	1,019	8.7	54.2	10.0	58.7

Table 12: 2002 cohort.

#### 9.3 Disclaimers

HM Revenue & Customs (HMRC) agrees that the figures and descriptions of results in the attached document may be published. This does not imply HMRC's acceptance of the validity of the methods used to obtain these figures, or of any analysis of the results.

Copyright of the statistical results may not be assigned. This work contains statistical data from HMRC which is Crown Copyright. The research datasets used may not exactly reproduce HMRC aggregates. The use of HMRC statistical data in this work does not imply the endorsement of HMRC in relation to the interpretation or analysis of the information.

The Student Loans Company (SLC) agrees that the figures and descriptions of results in

the attached document may be published. This does not imply SLC's acceptance of the validity of the methods used to obtain these figures, or of any analysis of the results.

Copyright of the statistical results may not be assigned. This work contains statistical data from SLC which is protected by Copyright, the ownership of which is retained by SLC. The research datasets used may not exactly reproduce SLC aggregates.

The use of SLC statistical data in this work does not imply the endorsement of SLC in relation to the interpretation or analysis of the information.